

```

C
C      CALCULATION BY CONTINUED FRACTION
C
30 A=1.0-P
   B=A+X+1.0
   TERM=0.0
   PN(1)=1.0
   PN(2)=X
   PN(3)=X+1.0
   PN(4)=X*B
   GIN=PN(3)/PN(4)
32 A=A+1.0
   B=B+2.0
   TERM=TERM+1.0
   AN=A*TERM
   DO 33 I=1,2
33 PN(I+4)=B*PN(I+2)-AN*PN(I)
   IF(PN(6).EQ.0.0) GO TO 35
   RN=PN(5)/PN(6)
   DIF=ABS(GIN-RN)
   IF(DIF.GT.ACU) GO TO 34
   IF(DIF.LE.ACU*RN) GO TO 42
34 GIN=RN
35 DO 36 I=1,4
36 PN(I)=PN(I+2)
   IF(ABS(PN(5)).LT.OFLO) GO TO 32
   DO 41 I=1,4
41 PN(I)=PN(I)/OFLO
   GO TO 32
42 GIN=I.0=FACTOR*GIN
C
50 GAMAIN=GIN
   RETURN
   END

```

## Algorithm AS 33

### Calculation of Hypergeometric Sample Sizes

By F. B. LEECH

*Rothamsted Experimental Station, Harpenden, Herts*

LANGUAGE

ASA Standard Fortran

PURPOSE

Solution, in  $n$ , of  $P = \{(N-m)!(N-n)!/\{N!(N-m-n)!\}$  for given values of  $P$ ,  $N$  and  $p = m/N$ .

The program was originally written in connection with a paper on innocuity testing of inactivated foot-and-mouth disease vaccines (Anderson *et al.*, 1970). In this context, a sample of  $n$  cm<sup>3</sup> from a vaccine batch of  $N$  cm<sup>3</sup> is tested for the presence of residual active virus. Batches giving any reaction are rejected and the theoretical problem is to select a size of sample such that when the test is negative, the amount, if any, of active virus in the batch must be small. In this context, the theory involves

only the first term of the hypergeometric distribution. In the equation above,  $P$  is the expected proportion of samples of  $n$  cm<sup>3</sup> negative to the test, when the batch of  $N$  cm<sup>3</sup> contains  $m$  infective cm<sup>3</sup>. The right-hand side is the first term of the hypergeometric distribution; corresponding equations, when 1, 2, ... cm<sup>3</sup> react to the test, involve the second and later terms of the distribution and the problem of solving them would be much more difficult.

The method uses the fact that  $P = \{1 - (n/N)\}^m$  is a very close approximation to the hypergeometric equation. Starting with trial values  $n = 0$  and  $NP$ , a new value of  $n$  is obtained by interpolation on the line  $\log(N - n) = a + b \log P$ . This process is repeated using the two most recent values of  $n$  until two successive estimates are the same. Within the range of parameters tested, the first interpolation was usually correct and never differed by more than one unit from the correct solution, which was reached on the next iteration. For many purposes, the simpler formula  $P = \{1 - (n/N)\}^m$ , which can readily be solved without a computer program, has all necessary accuracy.

#### STRUCTURE

##### *SUBROUTINE HYPERG (N,P,PP,K,IFault)*

##### *Formal parameters*

$N$  Integer input: as in the formula above.  
 $P$  Real input: the probability  $P$ .  
 $PP$  Real input: the proportion  $m/N$ .  
 $K$  Integer output: the solution.  
 $IFault$  Integer output: fault indicator.

Subroutine *HYPERG* has a function *FLOG* to calculation  $\log X!$ , using Stirling's approximation  $\frac{1}{2} \log(2\pi) + (X + \frac{1}{2}) \log X - X + 1/(12X)$ .

##### *Failure indication*

$IFault = 1$  if after 9 iterations the solution has not been reached;  
 $= 2$  if a proportion supplied by the external program not between 0 and 1;  
 $= 0$  otherwise.

##### *Restrictions*

None.

##### *Accuracy*

Results for small  $m$  can be checked arithmetically. These were found to be the nearest integral values to the non-integral exact solution.

#### REFERENCE

ANDERSON, E. C., CAPSTICK, P. B., MOWAT, G. N. and LEECH, F. B. (1970). A comparison between the use of cattle and a tissue culture system for safety testing of inactivated foot-and-mouth disease vaccines. *J. Hyg., Camb.*, **68**, 159-172.

```

SUBROUTINE HYPERG (N,PP,P,K,IFAU)
C
C     ALGORITHM AS 33 J.R.STATIST.SOC. C. (1970) VOL.19 NO.3
C
C     GIVEN N = NUMBER OF UNITS OF BATCH TO BE SAMPLED PP =
C     PROBABILITY OF POSITIVE UNITS BEING PRESENT, AND P= PROBABILITY
C     OF FAILING TO DETECT POSITIVE UNITS, FIND K, THE SIZE OF
C     SAMPLE TO BE DRAWN FROM THE BATCH
C
C     DOUBLE PRECISION A(10),B(10),BB,C,D,DBLE,X,FLOG,DLOG,DEXP
C
C     IFAU=0
C     IF((P.GT.1.0).OR.(P.LT.0.0)) GO TO 4
C     IF((PP.GT.1.0).OR.(PP.LT.0.0)) GO TO 4
C     X=N
C     D=DLOG(DBLE(P))
C     M=FLOAT(N)*PP+0.5
C     A(1)=0.00
C     B(1)=DLOG(X)
C     IF((P+PP).GT.1.) P=1.-PP
2    K=FLOAT(N)*P+0.5
C     I=N-K
C     X=I
C     C=DLOG(X)
C     N1=N-M
C     DO 3 L=2,10
C     N3=N1-K
C     A(L)=FLOG(N1)+FLOG(I)-FLOG(N)-FLOG((N3))
C     B(L)=C
C     BB=(B(L)-B(L-1))/(A(L)-A(L-1))
C     C=B(L)+BB*(D-A(L))
C     I=DEXP(C)+0.500
C     K1=N-I
C     IF(K.EQ.K1) RETURN
C     K=K1
3    CONTINUE
C     IFAU=1
C     SET DIAGNOSTIC FOR NON-CONVERGENCE
C     RETURN
4    IFAU=2
C     SET DIAGNOSTIC FOR DATA ERROR (PROPORTION NOT BETWEEN 0 AND 1)
C     RETURN
C     END
C
C     DOUBLE PRECISION FUNCTION FLOG(N)
C     GIVEN N, FIND LOG FACTORIAL N USING STIRLING'S APPROXIMATION
C     DOUBLE PRECISION F(6),X,DLOG
C     DATA F(1),F(2),F(3),F(4),F(5),F(6)/0.00,0.00,.693147800,
1    1.791759500,3.178053800,4.787491700/
C     IF(N.GT.5) GO TO 1
C     FLOG=F(N+1)
C     RETURN
1    X=N
C     FLOG=.9189385400+(X+.500)*DLOG(X)-X+1.000/(12.00*X)
C     RETURN
C     END

```