



Dose splitting increases selection for both target-site and non-target-site fungicide resistance – a modelling analysis

Journal:	<i>Plant Pathology</i>
Manuscript ID	Draft
Manuscript Type:	Original Article
Date Submitted by the Author:	n/a
Complete List of Authors:	Corkley, Isabel; Rothamsted Research, Net Zero and Resilient Farming; University of Reading, School of Agriculture, Policy and Development; ADAS, Sustainable Agricultural Systems Mikaberidze, Alexey; University of Reading, School of Agriculture, Policy and Development Paveley, Neil; ADAS, Sustainable Agricultural Systems van den Bosch, Frank; University of California Davis, Quantitative Biology & Epidemiology Group, Plant Pathology (Visiting Scholar); ADAS, Sustainable Agricultural Systems Shaw, Michael; University of Reading, School of Agriculture, Policy and Development Milne, Alice; Rothamsted Research, Net Zero and Resilient Farming
Topics:	control, chemical, epidemiology
Organisms:	fungi
Other Keywords:	fungicide resistance management, epidemiological model, septoria tritici blotch, non-target-site resistance, partial resistance, quantitative resistance

SCHOLARONE™
Manuscripts

Dose splitting increases selection for both target-site and non-target-site fungicide resistance – a modelling analysis

Isabel Corkley^{1,2,3}, Alexey Mikaberidze², Neil Paveley⁴, Frank van den Bosch^{5,6},

Michael W. Shaw², Alice E. Milne¹

¹*Net Zero and Resilient Farming, Rothamsted Research, Harpenden, Hertfordshire, AL5 2JQ, UK*

²*School of Agriculture, Policy and Development, University of Reading, Reading, UK*

³*Sustainable Agricultural Systems, ADAS, Wolverhampton, Unit 14 Newton Court, Pendeford Business Park, Wolverhampton, WV9 5HB, UK*

⁴*Sustainable Agricultural Systems, ADAS, High Mowthorpe, North Yorkshire, YO17 8BP, UK*

⁵*Quantitative Biology & Epidemiology Group, Plant Pathology Department (Visiting Scholar), University of California, Davis, One Shields Ave, Davis CA 95616, USA*

⁶*Sustainable Agricultural Systems, ADAS, Rosemaund, Hereford, HR1 3PG, UK*

Corresponding Author: I. Corkley; isabel.corkley@rothamsted.ac.uk

Keywords: fungicide resistance management; epidemiological model; septoria tritici blotch; non-target-site resistance; partial resistance; quantitative resistance

Abstract

Fungicide resistance management principles recommend that farmers avoid splitting the total dose applied of a fungicidal mode of action (MoA) across multiple applications per season ('dose splitting'). However, dose splitting may sometimes be needed to make another proven resistance management tactic - application in mixture with a different MoA - practically achievable, especially in cases where there are limited MoAs available for disease control. Variable effects of dose splitting on selection for resistance have been observed in field experiments, and its effect on selection for

partial resistance in fungal pathogens is not well studied. An improved understanding of whether the effect of dose splitting depends on fungicide properties and type of fungicide resistance is required. We developed a compartmental epidemiological model of septoria leaf blotch (STB) (*Zymoseptoria tritici*) to investigate the effect of dose splitting on selection for both complete and partial target-site and non-target-site resistance. To measure solely the effects of dose splitting, we restricted the analysis to solo fungicide application (solo use is not recommended in practice). Our results show variable effects of dose splitting: in general, it increased selection for both target-site and non-target-site resistance. Within the range of dose response parameters expected for commercial fungicides, dose splitting increased selection most for partial resistance mechanisms that result in a reduction in fungicide efficacy at low fungicide concentrations but not at high concentrations. We predict that dose splitting of a succinate dehydrogenase inhibitor (SDHI) fungicide (solo) will increase selection for target-site and non-target-site resistance by between 20-35%.

1. Introduction

The effectiveness of fungicides for control of plant diseases is threatened by the evolution of resistance (Corkley et al., 2022). The risk of resistance is particularly high for polycyclic foliar fungal pathogens, such as septoria tritici blotch (STB) (*Zymoseptoria tritici*) in wheat, grey mould (*Botrytis cinerea*) in many hosts, potato late blight (*Phytophthora infestans*), and net blotch (*Pyrenophora teres*) and powdery mildew (*Blumeria hordei*) diseases of barley. These pathogens have large population sizes and many generations per year, enabling rapid evolution of resistance (Grimmer et al., 2015; McDonald et al., 2022), and have the potential to cause large economic losses. Fungicide resistance management tactics include minimising the dose and number of applications, and applying in mixture with a different mode of action (MoA)

51 (Corkley et al., 2022; Elderfield et al., 2018; Mikaberidze et al., 2017; van den Berg et
52 al., 2016; van den Bosch et al., 2014a, 2014b). However, the number of effective MoA
53 available for use is increasingly restricted by regulation (especially of multi-site
54 fungicides) and resistance which has already evolved. This poses challenges for
55 implementation of current resistance management strategies.

56 Fungicides with a MoA affecting a single pathogen target site are at particular
57 risk of resistance development because a single point mutation affecting the target site
58 gene ('target-site resistance') may confer a large fitness advantage. Target-site
59 mutations may confer either complete or partial resistance. If a target-site mutation
60 substantially prevents fungicide binding, for example through a change in the shape
61 of the fungicide binding site, this can fully restore cellular or enzyme function and result
62 in a high level of complete resistance. For example, the G143A mutation prevents
63 quinone outside inhibitor (QoI) fungicides from binding to the cytochrome b
64 mitochondrial protein, restoring its function in respiration (Dorigan et al., 2023). Target-
65 site resistance may involve a single point mutation, or a combination of multiple
66 mutations on the target gene, each conferring partial resistance, but potentially leading
67 to highly resistant phenotypes in combination. For example, *Z. tritici* has accumulated
68 multiple mutations in the CYP51 gene, leading to gradually increasing levels of
69 resistance to demethylation inhibitor (DMI) fungicides (Cools & Fraaije, 2013; Hawkins
70 & Fraaije, 2021; Leroux & Walker, 2011). In addition to target-site mutations, other
71 mechanisms of fungicide resistance in pathogens include target-site overexpression,
72 and non-target-site resistance such as increased efflux, detoxification and alternative
73 metabolism (Dorigan et al., 2023; Hawkins & Fraaije, 2021; Hu & Chen, 2021). These
74 mechanisms may cause partially or highly resistant strains, especially in combination
75 with one another or with target-site resistance. Metabolic resistance pathways such as

efflux pumps are also implicated in multi-drug resistant fungal strains (Kretschmer et al., 2009; Omrane et al., 2017; Patry-Leclaire et al., 2023).

To predict the impact of fungicide resistance management tactics on selection, it is helpful to consider pathogen epidemics in terms of the *per capita* rate of increase or 'growth rate' (r) of each strain: a number which combines the repeating stages of lesion establishment, growth and sporulation into a single measure of the success of a strain at a given point in time. Pathogen strains with resistance to the action of a fungicide have higher growth rates in the presence of that fungicide than strains that are sensitive to the fungicide. The greater the difference in the *per capita* growth rates of resistant and sensitive strains, the faster the rate of selection for resistance (van den Bosch et al., 2014a). The impact of any given fungicide dose on the *per capita* growth rate of a pathogen strain can be represented in models by its effect on important parts of the pathogen life cycle, such as a reduction in the pathogen transmission rate. Assuming that the applied dose decays exponentially over time, it is possible to track the 'effective dose' remaining at any point in time. The impact of the fungicide on the pathogen life cycle is greatest at high effective doses, where the maximum effect is defined by an 'asymptote parameter', and the rate at which the effect decreases with reducing fungicide doses is defined by a 'curvature parameter'. The effect of resistance on the dose response to a fungicide may be observed either as a complete or partial reduction in the maximum effect of the fungicide on the pathogen growth rate even at very high effective doses, or as a reduction in the efficacy of lower effective doses of the fungicide. We will refer to these types of resistance as 'asymptote shift' and 'curvature shift' respectively, to reflect their effect on the fungicide dose response (Figure 1(a), 1(b)). Resistance resulting from an asymptote shift is sometimes referred to as 'qualitative' or 'type I' resistance, and

101 resistance resulting from a curvature shift as 'quantitative' or 'type II' resistance
102 (Elderfield, 2018; Mikaberidze et al., 2017; Taylor & Cunniffe, 2023a), but the definition
103 of these terms is not entirely consistent across the scientific literature.

104 Let us consider which resistance mechanisms are likely to lead to either a
105 partial asymptote shift or a curvature shift. Some fungicides bind competitively directly
106 to the enzyme active site: for example, DMI fungicides bind competitively to the CYP51
107 protein which catalyses a step in ergosterol biosynthesis (Hargrove et al., 2015),
108 occupying the P450 active site and preventing substrate binding. A target-site mutation
109 that causes a small to moderate reduction in the affinity of the enzyme for the fungicide
110 will reduce fungicide efficacy at low fungicide concentrations, but not at high fungicide
111 concentrations. This case is therefore best represented by a curvature shift. A
112 curvature shift will also be representative of other resistance mechanisms that reduce
113 fungicide efficacy at low fungicide concentrations but are overwhelmed by high
114 fungicide concentrations. These may include target-site overexpression and non-
115 target-site, metabolic resistance mechanisms such as increased expression of efflux
116 pumps and detoxification. A partial asymptote shift could result from a target-site
117 mutation that reduces the maximum effect at any dose rate of fungicides which bind
118 allosterically and non-competitively to an enzyme. These fungicides change the
119 structure of the enzyme in a way that inhibits enzyme function or reduces access or
120 binding of the substrate to the enzyme active site. An example is the cyanoacrylate
121 phenamacril which is used against a number of *Fusarium* species (Wollenberg et al.,
122 2020). The maximum effect of these fungicides could be partially reduced by a target-
123 site mutation which changes the shape of the enzyme-fungicide complex, partially
124 restoring enzyme function.

Multiple fungicide applications per year are often useful to avoid economically damaging epidemics of polycyclic foliar fungal pathogens such as *Z. tritici*. If the number of MoA available for programmes is limited, use of mixtures may require splitting the total dose of a fungicide across two or more applications, reducing the dose of each MoA per application but increasing the exposure time of the pathogen to each fungicide, with counteracting (but not necessarily equal) effects on selection for resistance. If resistance is evolving 'concurrently' to two or more MoA at the same time, this situation introduces complex trade-offs for resistance management. Whether 'splitting and mixing' is a good or a poor choice of strategy for management of concurrent evolution of resistance will depend on the balance between the effects of mixture and dose splitting on selection. However, variation in the effects of dose splitting is not well understood. van den Bosch et al. (2014a) hypothesise that dose splitting will, overall, increase selection for strains with an asymptote shift against a fungicide. They highlight several experimental studies that support this theory, but the effect of dose splitting on selection for partially resistant strains with a curvature shift has not been explicitly considered in previous modelling studies, to our knowledge. Field trials carried out between 2018 and 2020 to measure the effect of dose splitting on selection for SDH-mutants showed variable results (Paveley et al., 2020; Young et al., 2021). An improved understanding of how fungicide properties and type of resistance determine the effect of dose splitting on selection for resistant pathogen strains is needed to inform tactics for management of concurrent evolution of resistance.

To investigate the effect of dose splitting on selection, we developed a model of fungicide resistance evolution in *Z. tritici*. *Zymoseptoria tritici* is one of the most common, widespread and damaging pathogens affecting winter wheat crops in the UK

and worldwide. It has evolved resistance to Qols, DMIs and SDHIs (Cools & Fraaije, 2013; Dooley et al., 2016; Huf et al., 2018; Rehfus et al., 2018; Torriani et al., 2009), with a corresponding decline in disease control (Blake et al., 2018). The model simulates a typical UK epidemic of STB, describing the seasonal growth and senescence of the upper crop canopy of winter wheat under average temperature conditions in the UK, key processes in the pathogen life cycle (sporulation, infection and growth) and their interaction with fungicides. In the UK, initial infection of wheat crops by *Z. tritici* occurs in autumn or spring through airborne ascospores or by splash-dispersed conidia from wheat stubble. After penetrating the leaf stomata, the fungus develops slowly during a symptomless latent period, following which necrotic lesions form on the leaf surface. These produce asexual haploid pycnidiospores which spread to the upper leaf canopy through contact and rain splash, driving the majority of secondary infections within the growing season with the potential for rapid increases in disease severity (Ponomarenko et al., 2011; Suffert et al., 2011). STB is associated with a reduction in crop quality and yield losses of up to 50% if uncontrolled (Fones and Gurr, 2015).

Through model simulations, we compared the effects on selection for a resistant *Z. tritici* strain of applying a fungicide solo in either a single application at full label rate or in two applications, each at half the full label rate. It should be noted that use of solo MoA is not recommended in practice. However, restricting the analysis to dose splitting of a solo fungicide enabled us to measure solely the effects of dose splitting, rather than the combined effects of 'splitting and mixing', giving a clearer picture of the drivers in variation of the effects of dose splitting. We used the model to investigate how the effect of dose splitting on selection for resistance depends on: (a) fungicide properties (foliar concentration half-life; asymptote and curvature dose

response parameters for the sensitive strain); (b) the type of resistance (asymptote shift or curvature shift); and (c) the magnitude of the asymptote or curvature shift.

2. Materials and Methods

2.1 Model background and approach

We follow the approach of (Hobbelen et al., 2011b), modelling the leaf area index (LAI; a dimensionless measure of leaf density, defined as the total amount of one-sided leaf area of the canopy (m^2) per unit ground area (m^2)) and infection by *Z. tritici* pycnidiospores on the top three leaves of the wheat canopy only. Yield loss due to *Z. tritici* occurs due to a reduction in healthy leaf area duration (HAD) and the resulting loss of interception of photosynthetically active radiation (PAR) on the upper three leaves during grain-filling: the level of disease on the upper canopy is a good predictor of yield loss (Parker et al., 2004; Shaw & Royle, 1989). Fungicide applications targeted against *Z. tritici* are therefore mostly applied to the upper leaf canopy. Although there will be some fungicide exposure on lower leaves, previous modelling results suggest that it is on the upper leaf canopy that selection for resistance primarily occurs (van den Berg et al., 2013).

The dynamics of the epidemic in the model are driven by the growth and senescence of the crop, which determines the leaf area available for infection, and the effect of a fungicide on the pathogen life cycle over time. The leaf area can pass sequentially through healthy, latent (infected but not yet sporulating), infectious (sporulating) and post-infectious stages; healthy and latent leaf area may also senesce due to leaf age. The infectious leaf area generates new infections on healthy leaf area. The model simulates the LAI of both the latent and infectious stages of a sensitive strain and a resistant strain of *Z. tritici*.

Our model has the same functional form as one developed by Hobbelen *et al.* (2011a, b). However, the rate of senescence in that model was parameterised using data on spring barley (*Hordeum vulgare*) (Hobbelen *et al.*, 2011a), and the simulated timing of crop senescence could impact on model predictions of the effects of dose splitting on selection for resistant strains. We therefore re-parameterised the model (see Section 2.3) using a dataset of green leaf area index (GLAI) and *Z. tritici* infection of the top three leaves of wheat crops from 14 site-years (Milne *et al.*, 2003, described as 'Data set 1'; te Beest *et al.*, 2009).

2.2 Model equations

2.2.1 Growth and senescence of wheat leaf canopy

It is assumed that the growth rate of the total leaf area of the upper canopy is not affected by *Z. tritici* severity, so the total leaf area index (LAI) and uninfected healthy green leaf area index (GLAI) are tracked separately (Hobbelen *et al.*, 2011b). In the absence of disease the rates of change of the total LAI (A) and the total healthy GLAI (H) are given by:

$$\frac{dA}{dt} = \begin{cases} 0, & t < t_0 \\ \gamma(A_{\text{Max}} - A), & t > t_0 \end{cases} \quad (1)$$

$$\frac{dH}{dt} = \gamma(A_{\text{Max}} - A) - \beta(t) \quad (2)$$

$$\text{where } \beta(t) = \begin{cases} 0, & t < t_{\beta_0} \\ \tau \left(\frac{t - t_{\beta_0}}{t_{\beta_T} - t_{\beta_0}} \right) + \varphi e^{\omega(t_{\beta_T} - t)}, & t_{\beta_0} \leq t \leq t_{\beta_T} \end{cases} \quad (3)$$

where t_0 is the time at which leaf 3 emerges and growth of the upper canopy commences, A_{Max} is the maximum LAI, γ is the growth rate of the leaf area, $\beta(t)$ is the rate of senescence at time t , t_{β_0} is the time of onset of senescence, t_{β_T} is the time at which the canopy has fully senesced, and τ , φ and ω are coefficients controlling the rate at which senescence occurs in relation to the length of time after the onset of

senescence. Time is measured in degree days (base 0°C), 'zero-degree days' (see Section 2.3).

2.2.2 Infection of crop by *Zymoseptoria tritici*

The development of the STB epidemic is described in the model by tracking the LAI of latent and infectious lesions of the resistant and sensitive strains.

It is assumed that the epidemic on the upper leaves is initiated by an influx of spores from infectious lesions on lower leaves. The density of infectious lesions on lower leaves, C , diminishes over time at rate λ , as lower leaves senesce and infectious lesions on the lower leaves reach the end of the infectious period. The LAI of infectious lesions on lower leaves at time t , $C(t)$, is calculated as:

$$C(t) = C_0 e^{-\lambda t} \quad (4)$$

A fraction, $\theta_{\rho_{\text{Start}}}$, of the initial influx C from lower leaves is assumed to be spores of the resistant strain, with the sensitive strain fraction $\theta_{\sigma_{\text{Start}}} = 1 - \theta_{\rho_{\text{Start}}}$. It is assumed that $\theta_{\rho_{\text{Start}}}$ and $\theta_{\sigma_{\text{Start}}}$ are not affected by fungicide application after the start of the model simulation at GS31. The initial influx is denoted as C_{σ} and C_{ρ} for the sensitive and resistant strains respectively.

The influx of spores, C , and infectious LAI on the upper canopy, I , are converted into new latent lesions on the upper canopy, at transmission rate ε , i.e. the overall rate at which infectious lesion density is converted into new latent lesions on a given density of healthy leaf area. Latent lesions mature into infectious, sporulating lesions, at a rate δ , where $1/\delta$ is the average latent period. Infectious lesions die at a rate μ , where $1/\mu$ is the average infectious period. Leaf senescence affects latent LAI, but not infectious LAI as the leaf tissue is already killed by the necrotic process of lesions becoming infectious (Hobbelen et al., 2011b; Kema et al., 1996). The following set of equations track the area index of healthy (H), latently infected (L) and infectious (I)

leaf area over time, with L_ρ and L_σ denoting the area index of latent lesions and I_ρ and I_σ the infectious area index of the resistant and sensitive strains respectively:

$$\frac{dH}{dt} = \gamma(A_{\text{Max}} - A) - \beta(t)H - \varepsilon\left(\frac{H}{A}\right)(C_\sigma + C_\rho + I_\sigma + I_\rho) \quad (5)$$

$$\frac{dL_\sigma}{dt} = \varepsilon_\sigma\left(\frac{H}{A}\right)(C_\sigma + I_\sigma) - \delta L_\sigma - \beta(t)L_\sigma \quad (6)$$

$$\frac{dL_\rho}{dt} = \varepsilon_\rho\left(\frac{H}{A}\right)(C_\rho + I_\rho) - \delta L_\rho - \beta(t)L_\rho \quad (7)$$

$$\frac{dI_\sigma}{dt} = \delta_\sigma L_\sigma - \mu I_\sigma \quad (8)$$

$$\frac{dI_\rho}{dt} = \delta_\rho L_\rho - \mu I_\rho \quad (9)$$

The final fraction of the resistant strain in the population at crop senescence, $\theta_{\rho_{\text{End}}}$, is calculated as:

$$\theta_{\rho_{\text{End}}} = \frac{I_\rho(t_{\beta_T})}{I_\rho(t_{\beta_T}) + I_\sigma(t_{\beta_T})} \quad (10)$$

2.2.3 The effect of the fungicide on pathogen growth rate

Fungicide effects on the two strains of *Z. tritici* are simulated in the model through a dose-dependent reduction of pathogen life cycle parameters ε (transmission rate, Equations 6 and 7) and δ (the rate at which latent lesions are converted to sporulating lesions, Equations 8 and 9), slowing the rate of increase of the pathogen population. Single-site fungicides are assumed to reduce both the transmission rate and the rate of conversion of latent infections to sporulating lesions. The infectious period of sporulating lesions is assumed to be unaffected by fungicides.

The fungicide dose at time t , $D(t)$, is expressed as a proportion of the maximum permitted individual dose (as defined on the product label), D_{Max} , and decays exponentially over time at rate v :

$$D(t) = D_0 e^{-v(t-t^*)} \quad (11)$$

where D_0 is the applied dose and t^* is the time of application. $D(t)$ is the 'effective dose' referred to in Section 1.

The fungicide reduces the pathogen life cycle parameters ε and δ by a fraction $f(t)$, which changes over time depending on the remaining fungicide dose, $D(t)$. The dose response of $f(t)$ to $D(t)$ (Figures 1(a), 1(b)) is described by a combination of an asymptote parameter, q , which is the maximum fractional reduction of the pathogen life cycle parameter (i.e. at infinite fungicide dose), and a curvature parameter, k , which defines how quickly the fractional reduction declines from the asymptote as $D(t)$ decreases:

$$f_{\sigma}(t) = q_{\sigma}(1 - e^{-k_{\sigma}D(t)}) \quad (12)$$

$$f_{\rho}(t) = q_{\rho}(1 - e^{-k_{\rho}D(t)}) \quad (13)$$

The asymptote parameters are denoted as q_{σ} and q_{ρ} , the curvature parameters as k_{σ} and k_{ρ} , and the fractional reductions as $f_{\sigma}(t)$ and $f_{\rho}(t)$ for the sensitive and resistant strains respectively. Each pathogen life cycle parameter affected by the fungicide is multiplied by $(1 - f(t))$ to represent the effect of the fungicide on the growth rate of the pathogen population. For example, the transmission rate of the sensitive strain at time t , $\varepsilon_{\sigma}(t)$, is calculated as:

$$\varepsilon_{\sigma}(t) = \varepsilon_0(1 - f_{\sigma}(t)) = \varepsilon_0(1 - q_{\sigma}(1 - e^{-k_{\sigma}D(t)})) \quad (14)$$

where ε_0 is the transmission rate in the absence of fungicides. It is assumed that there are no fitness costs of resistance. If $f_{\sigma}(t) > f_{\rho}(t)$, the density of the resistant strain will increase faster than the density of the sensitive strain, leading to an increase in the resistant strain fraction of the *Z. tritici* population.

2.2.4 Types of fungicide resistance

We simulate two types of fungicide resistance based on the nature of the shift in sensitivity to the fungicide ('sensitivity shift'):

- Asymptote shift, ζ_q : parameter q is reduced relative to the sensitive strain.
- Curvature shift, ζ_k : parameter k is reduced relative to the sensitive strain.

We describe the level of sensitivity shift as a percentage. For example, a 50% asymptote shift means that $q_\rho = 0.5q_\sigma$. Partial resistance could take the form of either an asymptote shift or a curvature shift, or a combination of both. An asymptote shift means that the effect of any dose $D(t)$ against the resistant strain of the pathogen is reduced (Figure 1(a)). For a curvature shift, the instantaneous effect of a high dose of the fungicide may still be as potent, but at lower doses it is less effective against the resistant strain than against the sensitive strain (Figure 1(b)). The biological significance of asymptote and curvature shifts is discussed in Section 1.

A 100% asymptote and a 100% curvature shift are functionally identical: both represent strains that are completely resistant to the fungicide at any dose $D(t)$. Otherwise, for a given percentage sensitivity shift, an asymptote shift will result in a more highly resistant strain than the same level of curvature shift (as can be seen by comparing Figures 1(a) and 1(b)). The difference in the fractional reduction of the sensitive strain compared to the resistant strain, $f_\sigma(t) - f_\rho(t)$, is greatest at high fungicide dose $D(t)$ for asymptote shifts, and greatest at intermediate fungicide dose $D(t)$ for partial (<100%) curvature shifts (Figures 1(c), 1(d)).

2.2.5 Calculation of the selection coefficient

We used the selection coefficient, s , to compare the rate of selection for the resistant strain in each scenario simulated (Milgroom & Fry, 1988; van den Bosch et al., 2014a). The selection coefficient is defined as the difference in fitness between the resistant and sensitive strains due to the application of the fungicide, where fitness is measured by the per capita rate of increase, r , of a population:

$$s = r_\rho - r_\sigma \quad (15)$$

where r_ρ and r_σ are the average per capita rates of increase of the resistant and sensitive strains respectively over the course of the growing season. We calculate total selection between the start of the simulation, t_0 , and crop senescence, time t_{β_T} , denoting the total length of time simulated as T . Assuming exponential growth of the sensitive and resistant strains (in the absence of density dependence), the density of the sensitive strain and resistant strain at time t_{β_T} , denoted as $P_\sigma(t_{\beta_T})$ and $P_\rho(t_{\beta_T})$ respectively, can be calculated as:

$$P_\sigma(t_{\beta_T}) = P_\sigma(0)e^{r_\sigma T} \quad (16)$$

$$P_\rho(t_{\beta_T}) = P_\rho(0)e^{r_\rho T} \quad (17)$$

where $P_\sigma(0)$ and $P_\rho(0)$ are the initial densities of the sensitive and resistant strain respectively at the start of the simulation.

Rearrangement of equations (16) and (17) for r_σ and r_ρ , and substitution of equation (15) gives:

$$s = \frac{1}{T} \left(\ln \left(\frac{P_\rho(t_{\beta_T})P_\sigma(0)}{P_\rho(0)P_\sigma(t_{\beta_T})} \right) \right) \quad (18)$$

This can also be expressed in terms of the population fractions of the resistant and sensitive strains, θ_ρ and θ_σ , at the beginning of the simulation and the end of the growing season:

$$s = \frac{1}{T} \left(\ln \left(\frac{\theta_{\rho_{\text{End}}} \theta_{\sigma_{\text{Start}}}}{\theta_{\rho_{\text{Start}}} \theta_{\sigma_{\text{End}}}} \right) \right) \quad (19)$$

2.3 Model implementation and parameterisation

The model was implemented in MATLAB R2022b (The MathWorks Inc., 2022) using built-in function 'ode45' for the solution of the ordinary differential equations.

The model was parameterised using data on GLAI and *Z. tritici* infection over time from field trials of wheat crops grown with and without fungicide application, recorded over 14 site-years between 1993 and 1995 in England, United Kingdom, and

corresponding daily weather data from meteorological stations within one kilometre of the site (Milne et al., 2003, described as ‘Data set 1’; te Beest et al., 2009). We refer to data from these trials as ‘Dataset 1’. For each site-year, Dataset 1 includes data on four cultivars (Riband, Apollo, Slejpner and Haven), with four replicates per cultivar.

We chose to follow previous models (Elderfield et al. 2018; Hobbelen et al. 2011b; van den Berg et al. 2013) in parameterising the model on a zero-degree days scale. Weather data for the sites was used to calculate both the thermal time (degree days base 0°C) and photo-vernal-thermal time (base 1°C) since sowing (Milne et al., 2003; Weir et al., 1984) corresponding to each observation date. The photo-thermal-vernal time gave a more consistent profile for the timings of the upper canopy growth and senescence than thermal time (see Figure A.1.2 in Supporting Information A.1 for further details). Using linear regression, we derived a relationship between thermal time and photo-thermal-vernal time, t_{pvt} , and used this to convert t_{pvt} to the average thermal time in zero-degree days, t :

$$t = 1.204t_{pvt} + 778.6(20)$$

Dataset 1 was used to estimate the average number of zero-degree days per day, z .

We assumed that data from field plots that received a fungicide programme designed to provide full protection against disease (Milne et al., 2003) are representative of canopy growth in the absence of disease. We used these data to estimate the parameters controlling the growth and senescence of the wheat canopy: t_0 , t_{β_0} , t_{β_T} , A_{Max} , γ , τ , φ and ω (defined in Section 2.2.1). The mean GLAI of the top three leaves at each observation time point was calculated for each site-year from data from all four cultivars and replicates in Dataset 1. The parameters were fitted to data pooled from six site-years with maximum observed GLAI ranging from 3.76 to 4.90 (Cambridgeshire-1994, Devon-1994, Devon-1995, Kent-1995, Norfolk-1994,

Norfolk-1995), using least squares optimisation (lsqcurvefit, MATLAB 2022b; further details in Supporting Information A.1). Model zero-degree days were mapped to growth stages on Zadoks' scale (Zadoks et al., 1974), based on the fitted values of t_0 , t_{β_0} , t_{β_T} and the estimated phyllochron length (see Supporting Information A.1 for further details).

We estimated *Z. tritici* life cycle parameters δ , μ and λ (defined in Section 2.2.2) based on data from a literature search (Table 2). In combination with C_0 (Equation 4) and ε_0 (Equations 6, 7, 14), these parameters describe the infection of crop by *Z. tritici* in the absence of a fungicide. We estimated values for C_0 and ε_0 using data on STB epidemic progress (% severity) (Dataset 1) on untreated plots on which the maximum severity of the STB epidemic exceeded 5% and the maximum cumulative severity of yellow rust, brown rust and powdery mildew did not exceed 15%. Data from cultivars that were considered moderately resistant at the time the trials were carried out were used to estimate ε_0 . Data from six site-years (Devon-1994, Devon-1995, Hampshire-1995, Herefordshire-1994, Herefordshire-1995, Kent-1994) fitted these criteria. We fitted separate values of C_0 and ε_0 for each site-year-cultivar combination using least squares optimisation and calculated the average of these values (further details in Supporting Information A.1).

We used data from AHDB Fungicide Performance trials (AHDB, 2024a) on the observed dose response of STB severity to fluxapyroxad and isopyrazam from 2011-2012 (Dataset 2) to estimate indicative values of q_σ and k_σ for SDHI fungicides (see Supporting Information A.1 for further details), using an estimate of ν based on a literature search (Table 2).

2.4 Model simulations of dose splitting

392 We investigated the impact of dose splitting on selection for resistant strains with either
393 an asymptote shift or a curvature shift (either partial or complete resistance), for a
394 range of values of the fungicide parameters q_σ , k_σ and ν (Table 1). We compared
395 selection for the resistant strain following a single application of the fungicide at full
396 label rate, D_{Max} , at either growth stage 32 (GS32) or GS39, to selection for the resistant
397 strain following a 'split dose' application of $0.5D_{\text{Max}}$ at both GS32 and GS39. In all
398 simulations, the total dose applied to the upper leaf canopy, D_{Total} , was equal to D_{Max} .

399 The foliar concentration half-lives of fungicide products can be very variable
400 depending on the crop and environmental conditions (Fantke et al., 2014). We
401 simulated three values of ν (Table 1), equivalent to foliar half-lives of 3 days, 6 days
402 and 12 days; SDHI fungicides such as fluxapyroxad, penthiopyrad and fluopyram have
403 an average half-life of approximately 6 days (Fantke et al., 2014; He et al., 2016; Noh
404 et al., 2019). Figure 2 illustrates the effect of the decay rate on the simulated fungicide
405 dose $D(t)$ and fractional reduction $f(t)$ over time following single and split dose
406 applications.

407 We included very low and high values of parameters q_σ and k_σ in the analysis
408 to understand the extremes of the range of possible effects of dose-splitting. In
409 practice, these parameter values are unlikely in a commercially available fungicide:
410 fungicides with very low values of q_σ or k_σ would not be effective, whilst very high
411 values are more likely to be associated with an unacceptable toxicity profile. We
412 compared our results to those obtained using our fitted parameter values for SDHI
413 fungicides to understand the most likely range of effects of dose splitting on selection
414 for resistance to commercial fungicides.

415 We assumed that $\theta_\rho(0) = 0.01$, i.e. 1% of the inoculum initiating the epidemic
416 was the resistant *Z. tritici* strain, whilst the remaining 99% of the population was

sensitive to the fungicide. The simulations were run for a single growing season from the start of the leaf growth of the upper canopy, t_0 , to complete canopy senescence, t_{β_T} . For each combination of parameter values simulated, the selection coefficient for the resistant strain, s , was calculated (Equation 19). The percentage change in the selection coefficient due to dose splitting, η , was then calculated as:

$$\eta = 100 \times \frac{(s_{\text{Split}} - s_{\text{Single}})}{s_{\text{Single}}} \quad (21)$$

where s_{Single} is the selection coefficient for a single application at D_{Total} and s_{Split} is the selection coefficient for the resistant strain for a split dose application.

3. Results

3.1 Model parameterisation

The fitted model parameters are summarised in Table 2. The model fit to observed GLAI in the absence of disease was good (Figure 3(a); $n=76$, $R^2 = 76.9\%$, $\text{RMSE} = 0.76$). For the cultivar-site-year combinations used to fit ε_0 , the transmission rate in the absence of fungicide, the overall fit to observed disease severity progress was excellent ($n=293$, $R^2 = 88.4\%$, $\text{RMSE} = 2.8\%$); fitted values of ε_0 ranged from 0.0136 to 0.0364, with a mean value of 0.0211. In the absence of a fungicide, the model predicts STB severity of 9.5% (Figure 3(b)) at GS75 (medium milk), which is approximately equivalent to the expected average severity on a cultivar with an AHDB resistance rating of 6 (AHDB, 2024b).

3.2 Effect of dose splitting on selection for fungicide resistance

For the range of parameter values simulated (Table 1), we show results for both the overall magnitude of selection, measured by the selection coefficient s (Section 2.2.5), and the percentage change in selection due to dose splitting, η (Equation 21). When describing the baseline level of efficacy of a fungicide in Sections 3.2.1 and 3.2.2, we refer to the dose response against the sensitive strain, notated as q_σ and

442 k_{σ} for the asymptote and curvature parameter respectively. For a resistant strain with
443 an asymptote shift, $\zeta_q > 0$ but no curvature shift i.e. $\zeta_k = 0$, note that $k_{\rho} = k_{\sigma}$. For a
444 resistant strain with a curvature shift $\zeta_k > 0$ but no asymptote shift, $q_{\rho} = q_{\sigma}$.

445 3.2.1 Magnitude of selection

446 The magnitude of selection for fungicide resistance, measured by the selection
447 coefficient s , increased for both single and split dose fungicide applications with
448 increasing values of the asymptote parameter, q_{σ} , curvature parameter, k_{σ} , asymptote
449 shift, ζ_q or curvature shift, ζ_k , and with decreasing values of the decay rate, ν (Figure
450 4). This means that a strain with resistance against a highly effective fungicide (with
451 high values of q_{σ} , k_{σ} and a relatively low value of ν) would spread more quickly if the
452 fungicide was applied, compared to a strain with resistance against a fungicide with
453 lower efficacy. The greater the effect of a fungicide on the growth rate of the sensitive
454 strain, the greater the maximum magnitude of the cumulative difference in growth rates
455 between the resistant and sensitive strains when the fungicide is applied. More highly
456 resistant strains (higher values of ζ_q or ζ_k) will also spread more quickly, as they have
457 higher growth rates in the presence of a fungicide relative to the sensitive strain.

458 As noted in Section 2.2.4, either a 100% asymptote shift or 100% curvature
459 shift leads to a strain that is completely resistant to the fungicide at any dose $D(t)$, and
460 an identical value of s for a given combination of q_{σ}, k_{σ} and ν . For a given sensitivity
461 shift percentage less than 100% (e.g. 50% or 90%), s is higher for an asymptote shift
462 than for the same level of curvature shift, as the asymptote shift corresponds in a more
463 highly resistant strain, leading to a greater cumulative difference in growth rates
464 between the resistant and sensitive strain when fungicide is applied.

465 For partial and complete asymptote shifts, s was consistently higher for split
466 dose applications than for single applications.

3.2.2 Effect of dose splitting on selection for resistance, η

The values of the asymptote parameter, q_σ , and asymptote shift, ζ_q , have very little impact on the percentage change in the selection coefficient s (η in Equation 21) as a result of dose splitting (Figure 5). q_σ also has very little impact on η for a curvature shift (Figure A.2.1, Supporting Information A.2). This is because q_σ and ζ_q do not affect the length of time for which there is a difference in the level of control exerted by single and split dose applications. The curvature parameter, k_σ , and the decay rate, ν , together control the value of η , in combination with the curvature shift, ζ_k , where relevant (Figure 6).

For any asymptote shift, dose splitting increased selection for resistance. The value of η for an asymptote shift varied from <5% to 40%, depending on the values of k_σ and ν (Figure 6(a)-(c)). Our results suggest that splitting the dose of a solo SDHI across two applications rather than making a single application at full dose rate could increase selection for a strain with an asymptote shift to the SDHI by approximately 20%.

For curvature shifts, η varied from -20% to 80% (Figure 6(d)-(f)), indicating that dose splitting can reduce selection for partially resistant strains in some cases, but in other cases it may lead to a large increase in selection for resistance, dependent on the values of k_σ , ν and ζ_k . The value of η increased with the curvature parameter, k_σ , reaching an asymptote at high values of k_σ when the fungicide half-life was short (Figure 6(d)). For longer fungicide half-lives, the value of η initially increased with k_σ to a maximum, then decreased at very large values of k_σ (Figure 6(f)). For larger curvature shifts, ζ_k , the η -values approach the curves for asymptote shifts (Figure 6(a)-(c)). For smaller curvature shifts, $\zeta_k < 50\%$, η initially increased with k_σ , to a maximum at approximately $5 \leq k_\sigma \leq 10$, and then decreased again for

larger values of k_σ . For small curvature shifts, ζ_k , large curvature parameters, k_σ , and longer fungicide half-lives, η approached zero or even became negative. Our results suggest that dose splitting of a solo SDHI application would increase selection for a strain with a curvature shift to the SDHI by approximately 20-35%, with smaller curvature shifts falling towards the upper end of this range.

Dose splitting will increase selection for resistance if it leads to a larger difference in the growth rates of the sensitive strain and resistant strain for a longer time than a single application, i.e. if it increases the overall sum of the differences in fractional reduction, $\sum_{t=0}^T (f_\sigma(t) - f_\rho(t))$. For an asymptote shift, the maximum difference in the growth rates of the sensitive strain and the resistant strain occurs at high fungicide doses, $D(t)$, for which the fractional reduction $f_\sigma(t)$ is close to the maximum (as defined by the asymptote q_σ) (Figure 1(c)). For a curvature shift, dose response curves for sensitive and resistant strains converge at high values of $D(t)$. The maximum difference in the fractional reduction and resulting growth rates of the sensitive strain and a resistant strain with a curvature shift occurs at intermediate fungicide dose $D(t)$ (Figure 1(d)). As discussed by Taylor & Cunniffe (2023b), the effect of dose-response convergence on selection must be considered not only at the applied dose, but across the full time span of fungicide decay. Dose splitting increases the length of time that the pathogen is exposed to intermediate fungicide doses, which therefore increases $\sum_{t=0}^T (f_\sigma(t) - f_\rho(t))$. The results in Figure 6 can be understood by considering how the values of k_σ , ν and ζ_k affect the size and duration of the difference in the growth rates of the sensitive and resistant strain, for single and split dose applications.

Effect of decay rate, ν

For both asymptote shifts and curvature shifts, η was higher for larger values of ν (Figure 6). If the decay rate is high, the effect of a single application dissipates quickly, so a split dose application is likely to double the exposure time. If the decay rate is low, the effect of a single application at full dose rate will last for longer, so there is less difference in exposure time compared to the split dose application.

Why does η increase with k_σ for asymptote shifts?

For small values of the curvature parameter k_σ (approx. <4), the maximum reduction of the sensitive strain life cycle parameters is only achieved at a high fungicide dose, $D(t)$, and the fractional reduction reduces quickly as $D(t)$ decreases (Figure A.2.2(a), Supporting Information A.2). Therefore, the higher maximum dose applied in the single application initially achieves a much higher fractional reduction than the split dose application. Larger corresponding differences in the growth rates of the resistant and sensitive strain partially counterbalance the increased selection from the increased exposure time in the split dose application. The rate of selection from either a single or split dose application is therefore relatively similar for small values of k_σ , resulting in small values of η .

As k_σ increases, the fractional reduction remains close to the maximum fractional reduction even at lower fungicide doses $\leq 0.5 D_{Max}$, so at lower values of $D(t)$, differences in the growth rates of the resistant and sensitive strain are similar to the difference at the full dose rate (Figure A.2.2(b), Supporting Information A.2). The effect of the increased exposure time from the split dose therefore dominates at higher values of k_σ , resulting in higher values of η .

Why does η exhibit a maximum vs. k_σ for asymptote shifts when ν is low?

If k_σ is large and ν is low, the effect of a single application persists close to the maximum fractional reduction for a long time (Figure 2(f); Figure A.2.2(c), Supporting

Information A.2), which shifts the point at which there is a large difference in the fractional reduction from the single application and the split dose application later in the season. Since canopy senescence begins to restrict the growth rates of both the resistant and sensitive strains later in the season, the value of η is reduced relative to the maximum at intermediate values of k_σ and lower values of ν . However, the effect of dose splitting may still be larger than for small values of k_σ .

Why does η increase with k_σ more for curvature shifts than for asymptote shifts?

As k_σ increases, the dose response curve for the sensitive strain becomes more steeply curved, resulting in a decrease in the fungicide dose $D(t)$ at which the difference $f_\sigma(t) - f_\rho(t)$ is maximised for a curvature shift. The larger the value of k_σ and the smaller the value of ζ_k , the lower the dose $D(t)$ at which the difference $f_\sigma(t) - f_\rho(t)$ is maximised (Figure 1; Figure A.2.2(d)-(f), Supporting Information A.2), as resistant strains with a small curvature shift are still well controlled at high fungicide doses.

For very small values of k_σ , the maximum difference in growth rates occurs at higher values of $D(t) > 0.5D_{Max}$, which may not be reached using a split dose application. The maximum difference in growth rates is reached by the higher dose rate of the single application, partially counterbalancing the increased exposure time from the split dose application. Therefore η is small for small values of k_σ for a curvature shift. For larger values of k_σ , the maximum difference in growth rates occurs at values of $D(t) < 0.5D_{Max}$. A split dose application keeps $D(t)$ close to the level that maximises $f_\sigma(t) - f_\rho(t)$ for longer. In combination with the effect of increased exposure time, a split dose application increases selection more for strains with a curvature shift than for strains with an asymptote shift for intermediate values of k_σ .

565 *Why does η become negative for small curvature shifts, large values of k_σ and small*
566 *values of ν ?*

567 If k_σ is large and ζ_k is small, the maximum difference in growth rates occurs at very
568 small values of $D(t) < 0.1D_{Max}$ (Figure A.2.2(f), Supporting Information A.2). If the
569 decay rate, ν , is also small, low values of $D(t)$ are not reached for a split dose
570 application until late in the season, when canopy senescence restricts the growth rates
571 of both the resistant and sensitive strains, leading to low or even negative values of η
572 for large values of k_σ combined with small values of ν and small values of ζ_k .

573 It is important to note that our results do not suggest that there would be no
574 selection for resistance in cases where η was close to 0 or even negative: on the
575 contrary, selection for resistance will usually be strong in cases with large values of k_σ
576 and small values of ν (Figure 4), as resistance against a very effective fungicide gives
577 a strong fitness advantage. However, in these cases dose splitting may have little
578 effect on the strength of selection for resistance, or may even slightly decrease
579 selection relative to a single application.

580 **4. Discussion**

581 Dose splitting is likely to increase selection for both target-site and non-target-site
582 resistance. Our results suggest that the percentage increase in selection due to dose
583 splitting, η , is likely to be particularly large for resistance mechanisms that cause a
584 curvature shift, where the effect of the fungicide is reduced at lower concentrations but
585 not at high concentrations. These mechanisms could include non-target-site
586 resistance, target-site overexpression, and target-site mutations that affect fungicide
587 competitive binding rates. Our results also support the hypothesis of van den Bosch
588 et al. (2014a) that dose splitting will increase selection for target-site mutations that
589 cause an asymptote shift.

We show that the effects of dose splitting can be very variable for both target-site and non-target-site resistance. The largest increases in selection due to dose splitting are likely to occur for fungicides with a steeply curved dose response curve (i.e. high values of k_{σ}) and a relatively short half-life (i.e. high values of the decay rate, ν). In these cases, dose splitting should be considered high-risk for both target-site and non-target-site resistance. Our analysis focused on dose splitting of a solo MoA, whereas resistance management guidelines recommend application in mixture with other MoA; mixture may reduce selection for resistance and change the measured effects of dose splitting (Young et al., 2021). Where use of mixture requires 'splitting and mixing' due to limited numbers of effective MoAs for use in disease control, careful choice of mixture partners will be needed for fungicides for which dose splitting is high-risk for resistance evolution.

We found a small range of parameter values – fungicides with a large curvature parameter and a low decay rate – for which dose splitting could reduce selection for a resistant strain with a small curvature shift. However, these parameter values are relatively unlikely for a commercial fungicide, unless a high level of persistence could be achieved without associated environmental toxicity that would prevent regulatory approval. We used SDHI fungicides as an example of a commercial MoA currently available to growers. Our results suggest that dose splitting of an SDHI fungicide applied solo will increase selection for resistance by 20-35%.

Our results suggest that variability in fungicide decay rates between years and sites due to differing environmental conditions is likely to contribute to the variable selection for SDH-mutants observed in field experiments on dose splitting (Paveley et al. 2020; Young et al. 2021). We modelled the effect of a 4-fold change in fungicide half-life, which is well within the maximum range observed in field conditions (Fantke

et al., 2014). Our results suggest that for a fungicide with $k_{\sigma} = 10$, the variation in decay rates could account for the variation in the percentage effect of dose splitting on selection, η , in the range 10-40% for an asymptote shift, or 0-70% for a curvature shift (Figures 6(b), 6(e)). The statistical power of field trials to detect the lower end of this range may be limited due to experimental noise, but our results confirm that dose splitting tends to increase selection for resistance.

There is a strong covariance between the fitted values of k_{σ} , q_{σ} , and ν for the SDHI fungicide, increasing uncertainty in the estimation of these parameters and the consequences of dose splitting. We also assumed that k_{σ} and q_{σ} were the same for the fractional reduction of the transmission rate and the rate of conversion from latent to infectious leaf tissue. Measures of fungicide foliar half-life for each trial, and laboratory investigation of the effects of different fungicide dose rates on life cycle parameters such as latent period, could provide valuable additional evidence to inform these parameter values.

In our study we assumed negligible fitness costs of fungicide resistance, which is often the case (Hawkins & Fraaije, 2018; Mikaberidze & McDonald, 2015). However, fitness costs may sometimes suppress the growth rate of the resistant strain to a level below the growth rate of the sensitive strain. This can occur in the absence of fungicide, at low fungicide doses for an asymptote shift (Mikaberidze et al., 2017), or at high fungicide doses for resistant strains with a small curvature shift. Fitness costs have been reported for some target-site and non-target-site mutations; conversely, resistant strains can also have increased virulence relative to wild-type strains (Dorigan et al., 2023).

We did not explicitly model polygenic resistance, where resistance is conferred by multiple genes and the degree of resistance can build up gradually over time as

640 resistance mutations accumulate. At the population level, this process leads to a
641 continuous distribution of resistance phenotypes across strains, with the average
642 levels of resistance increasing over time as selection for resistance continues (Shaw,
643 1989; Taylor & Cunniffe, 2023a). The difference between the dose response curves
644 of partially-resistant strains may be analogous to a small curvature shift in our model,
645 meaning that dose splitting could strongly increase the rate of selection for polygenic
646 resistance.

647 The variable effect of dose splitting complicates management of resistance
648 evolving 'concurrently' to two or more MoA at the same time. Use of mixtures may
649 require splitting the total dose of a fungicide across two or more applications, due to a
650 limited number of MoA available. The balance between the effects of mixture and dose
651 splitting on selection for resistance will change depending on fungicide properties and
652 resistance type and strength, and the optimal strategy to slow evolution of resistance
653 to one fungicide may not be the optimal strategy for another fungicide. The efficacy of
654 the fungicide programme also needs to be considered and, where relevant, the effects
655 of sexual reproduction of the pathogen.

656 Previous modelling studies found that if it is necessary to combine two high-risk
657 fungicides in a programme, mixture rather than alternation or concurrent use will
658 generally present the best strategy to maximise the length of time that effective
659 disease control can be maintained (Elderfield, 2018; Hobbelen et al., 2013). However,
660 Elderfield (2017) found that alternation may be a better strategy against strains with
661 a small curvature shift. Experimental evolution *in vitro* on sensitive isolates of *Z. tritici*
662 using mixtures of high-risk fungicides showed that the success of mixture in delaying
663 resistance depended strongly on the mixture components, and some reduced-dose
664 mixtures selected for generalist, multi-drug resistance (Ballu et al., 2021). These

results may be explained by our finding that dose splitting increases selection more for strains with a small curvature shift – representative of non-target-site resistance – than for strains with an asymptote shift.

Since the balance between the effects of mixture and dose splitting on selection for resistance will differ for asymptote and curvature shifts, this could introduce trade-offs between tactics to reduce selection for large, target-site, asymptote shifts and alternative tactics to limit incrementally increasing levels of resistance due to mechanisms that cause a curvature shift. These trade-offs appear to occur in weed management, where use of herbicide mixtures is associated with lower prevalence of target-site resistance, but higher prevalence of metabolic resistance (Comont et al., 2020). Fungicide resistance management strategies have tended to focus on large asymptote shifts associated with target-site mutations, as these can lead to a rapid loss of fungicide efficacy, for example as experienced in QoI fungicides for multiple pathogens (Grimmer et al., 2015). Due to their large effects, target-site mutations that result in an asymptote shift are more likely to be quickly identified and studied than individual non-target-site resistance mechanisms which may be overlooked due to the small effects of each gene (Hu and Chen, 2021). However, in combination with target-site resistance, non-target-site mechanisms may contribute to highly resistant MDR strains (Omrane et al., 2017). Synergistic interactions between resistance mechanisms could enhance the overall impact of non-target site resistance: for example, increased efflux reduces the cellular fungicide concentration and could therefore increase the effect of a target-site mutation that causes a partial curvature shift. Wherever possible, tactics should be chosen for their effectiveness against both target-site and non-target-site resistance.

5. Acknowledgements

This research was funded by AHDB (project 21120062). Rothamsted Research receives strategic funding from the Biotechnology and Biological Sciences Research Council of the United Kingdom. AEM acknowledges support from the Growing Health Institute Strategic Programme (BBS/E/RH/230003C).

6. Data availability statement

Dataset 1: Data sharing is not applicable to this dataset as no new data were created or analysed in this study.

Dataset 2: These data are available from the Agriculture and Horticulture Development Board (AHDB). Restrictions apply to the availability of these data, which were used under license for this study. A summarized version of the data used is available at <https://ahdb.org.uk/knowledge-library/a-guide-to-fungicide-performance-in-wheat-barley-and-oilseed-rape>, and in Supporting Information A.1 (Figure A.1.5).

7. References

AHDB (2024a) A guide to fungicide performance in wheat, barley and oilseed rape.

Available at: <https://ahdb.org.uk/knowledge-library/a-guide-to-fungicide-performance-in-wheat-barley-and-oilseed-rape> [Accessed: 1 August 2024].

AHDB (2024b) Recommended Lists disease ratings. Available at:

<https://ahdb.org.uk/recommended-lists-disease-ratings> [Accessed: 1 August 2024].

Ballu, A., Deredec, A., Walker, A.-S., & Carpentier, F. (2021) Are Efficient-Dose

Mixtures a Solution to Reduce Fungicide Load and Delay Evolution of

Resistance? An Experimental Evolutionary Approach. *Microorganisms* 9, 2324.

<https://doi.org/10.3390/microorganisms9112324>

- 713 Blake, J.J., Gosling, P., Fraaije, B.A., Burnett, F.J., Knight, S.M., Kildea, S. et al.
714 (2018) Changes in field dose–response curves for demethylation inhibitor (DMI)
715 and quinone outside inhibitor (QoI) fungicides against *Zymoseptoria tritici*,
716 related to laboratory sensitivity phenotyping and genotyping assays. Pest
717 Manag Sci 74, 302–313. <https://doi.org/10.1002/ps.4725>
- 718 Boixel, A.-L. (2020). Environmental heterogeneity, a driver of adaptation to
719 temperature in foliar plant pathogen populations? PhD Thesis. Université Paris-
720 Saclay. Available at: <https://pastel.hal.science/tel-03202132/document>
721 [Accessed: 1 August 2024].
- 722 Comont, D., Lowe, C., Hull, R., Crook, L., Hicks, H.L., Onkokesung, N., et al. (2020)
723 Evolution of generalist resistance to herbicide mixtures reveals a trade-off in
724 resistance management. Nat Commun 11, 3086.
725 <https://doi.org/10.1038/s41467-020-16896-0>
- 726 Cools, H.J. & Fraaije, B.A. (2013) Update on mechanisms of azole resistance in
727 *Mycosphaerella graminicola* and implications for future control. Pest Manag Sci
728 69, 150–155. <https://doi.org/10.1002/ps.3348>
- 729 Corkley, I., Fraaije, B. & Hawkins, N. (2022) Fungicide resistance management:
730 Maximizing the effective life of plant protection products. Plant Pathol 71.
731 <https://doi.org/10.1111/ppa.13467>
- 732 Dooley, H., Shaw, M.W., Mehenni-Ciz, J., Spink, J. & Kildea, S. (2016) Detection of
733 *Zymoseptoria tritici* SDHI-insensitive field isolates carrying the *SdhC* -H152R
734 and *SdhD* -R47W substitutions. Pest Manag Sci 72, 2203–2207.
735 <https://doi.org/10.1002/ps.4269>
- 736 Dorigan, A.F., Moreira, S.I., da Silva Costa Guimarães, S., Cruz-Magalhães, V. &
737 Alves, E. (2023) Target and non-target site mechanisms of fungicide resistance

- 738 and their implications for the management of crop pathogens. *Pest Manag Sci*
739 79, 4731–4753. <https://doi.org/10.1002/ps.7726>
- 740 Elderfield, J.A.D. (2018) Using epidemiological principles and mathematical models
741 to understand fungicide resistance evolution. PhD thesis, University of
742 Cambridge. <https://doi.org/10.17863/CAM.22236>
- 743 Elderfield, J.A.D., Lopez-Ruiz, F.J., van den Bosch, F. & Cuniffe, N.J. (2018) Using
744 Epidemiological Principles to Explain Fungicide Resistance Management
745 Tactics: Why do Mixtures Outperform Alternations? *Phytopathology* 108, 803–
746 817. <https://doi.org/10.1094/PHYTO-08-17-0277-R>
- 747 Eyal, Z. (1971) The kinetics of pycnospore liberation in *Septoria tritici*. *Canadian*
748 *Journal of Botany* 49, 1095–1099. <https://doi.org/10.1139/b71-157>
- 749 Fantke, P., Gillespie, B.W., Juraske, R. & Joliet, O. (2014). Estimating Half-Lives for
750 Pesticide Dissipation from Plants. *Environ Sci Technol* 48, 8588–8602.
751 <https://doi.org/10.1021/es500434p>
- 752 Fones, H. & Gurr, S. (2015). The impact of *Septoria tritici* Blotch disease on wheat:
753 An EU perspective. *Fungal Genetics and Biology* 79, 3–7.
754 <https://doi.org/10.1016/j.fgb.2015.04.004>
- 755 Grimmer, M.K., van den Bosch, F., Powers, S.J. & Paveley, N.D. (2015) Fungicide
756 resistance risk assessment based on traits associated with the rate of pathogen
757 evolution. *Pest Manag Sci* 71, 207–215. <https://doi.org/10.1002/ps.3781>
- 758 Hargrove, T.Y., Wawrzak, Z., Lamb, D.C., Guengerich, F.P. & Lepesheva, G.I.
759 (2015) Structure-Functional Characterization of Cytochrome P450 Sterol 14 α -
760 Demethylase (CYP51B) from *Aspergillus fumigatus* and Molecular Basis for the
761 Development of Antifungal Drugs. *Journal of Biological Chemistry* 290, 23916–
762 23934. <https://doi.org/10.1074/jbc.M115.677310>

- 763 Hawkins, N.J. & Fraaije, B.A. (2021) Contrasting levels of genetic predictability in the
764 evolution of resistance to major classes of fungicides. *Mol Ecol* 30, 5318–5327.
765 <https://doi.org/10.1111/mec.15877>
- 766 Hawkins, N. J. & Fraaije, B. A. (2018). Fitness Penalties in the Evolution of Fungicide
767 Resistance. *Annual Review of Phytopathology*, 56(1), 339–360.
768 <https://doi.org/10.1146/annurev-phyto-080417-050012>
- 769 He, M., Jia, C., Zhao, E., Chen, L., Yu, P., Jing, J. et al. (2016) Concentrations and
770 dissipation of difenoconazole and fluxapyroxad residues in apples and soil,
771 determined by ultrahigh-performance liquid chromatography electrospray
772 ionization tandem mass spectrometry. *Environmental Science and Pollution*
773 *Research* 23, 5618–5626. <https://doi.org/10.1007/s11356-015-5750-6>
- 774 Hobbelen, P.H.F., Paveley, N.D., Fraaije, B.A., Lucas, J.A. & van den Bosch, F.
775 (2011a) Derivation and testing of a model to predict selection for fungicide
776 resistance. *Plant Pathol* 60, 304–313. [https://doi.org/10.1111/j.1365-](https://doi.org/10.1111/j.1365-3059.2010.02380.x)
777 [3059.2010.02380.x](https://doi.org/10.1111/j.1365-3059.2010.02380.x)
- 778 Hobbelen, P.H.F., Paveley, N.D., Oliver, R.P. & van den Bosch, F. (2013). The
779 Usefulness of Fungicide Mixtures and Alternation for Delaying the Selection for
780 Resistance in Populations of *Mycosphaerella graminicola* on Winter Wheat: A
781 Modeling Analysis. *Phytopathology* 103, 690–707.
782 <https://doi.org/10.1094/PHYTO-06-12-0142-R>
- 783 Hobbelen, P.H.F., Paveley, N.D. & van den Bosch, F. (2011b). Delaying selection for
784 fungicide insensitivity by mixing fungicides at a low and high risk of resistance
785 development: A modeling analysis. *Phytopathology* 101, 1224–1233.
786 <https://doi.org/10.1094/PHYTO-10-10-0290>

- 787 Hu, M. & Chen, S. (2021) Non-Target Site Mechanisms of Fungicide Resistance in
788 Crop Pathogens: A Review. *Microorganisms* 9, 502.
789 <https://doi.org/10.3390/microorganisms9030502>
- 790 Huf, A., Rehfus, A., Lorenz, K.H., Bryson, R., Voegelé, R.T. & Stämmler, G. (2018)
791 Proposal for a new nomenclature for *CYP51* haplotypes in *Zymoseptoria tritici*
792 and analysis of their distribution in Europe. *Plant Pathol* 67, 1706–1712.
793 <https://doi.org/10.1111/ppa.12891>
- 794 Kema, G.H.J., Yu, D., Rijkenberg, F.H.J., Shaw, M.W. & Baayen, R.P. (1996)
795 Histology of the Pathogenesis of *Mycosphaerella graminicola* in Wheat.
796 *Phytopathology* 86, 777. <https://doi.org/10.1094/Phyto-86-777>
- 797 Kretschmer, M., Leroch, M., Mosbach, A., Walker, A.-S., Fillinger, S., Mernke, D. et
798 al. (2009) Fungicide-Driven Evolution and Molecular Basis of Multidrug
799 Resistance in Field Populations of the Grey Mould Fungus *Botrytis cinerea*.
800 *PLoS Pathog* 5, e1000696. <https://doi.org/10.1371/journal.ppat.1000696>
- 801 Leroux, P. & Walker, A. (2011) Multiple mechanisms account for resistance to sterol
802 14 α -demethylation inhibitors in field isolates of *Mycosphaerella graminicola*.
803 *Pest Manag Sci* 67, 44–59. <https://doi.org/10.1002/ps.2028>
- 804 McDonald, B.A., Suffert, F., Bernasconi, A. & Mikaberidze, A. (2022) How large and
805 diverse are field populations of fungal plant pathogens? The case of
806 *Zymoseptoria tritici*. *Evol Appl* 15, 1360–1373. <https://doi.org/10.1111/eva.13434>
- 807 Mikaberidze A. & McDonald, B. A. (2015) Fitness cost of fungicide resistance -
808 impact on management. Invited chapter in the book “Fungicide Resistance in
809 Plant Pathogens: Principles and a Guide to Practical Management”, p. 77-89,
810 Springer Japan, eds. Hideo Ishii & Derek Hollomon. https://doi.org/10.1007/978-4-431-55642-8_6.
811

- 812 Mikaberidze, A., Paveley, N., Bonhoeffer, S. & van den Bosch, F. (2017) Emergence
813 of Resistance to Fungicides: The Role of Fungicide Dose. *Phytopathology* 107,
814 545–560. <https://doi.org/10.1094/PHYTO-08-16-0297-R>
- 815 Milgroom, M.G. & Fry, W.E. (1988). A Simulation Analysis of the Epidemiological
816 Principles for Fungicide Resistance Management in Pathogen Populations.
817 *Phytopathology* 78, 565. <https://doi.org/10.1094/Phyto-78-565>
- 818 Milne, A., Paveley, N., Audsley, E. & Livermore, P. (2003) A wheat canopy model for
819 use in disease management decision support systems. *Annals of Applied*
820 *Biology* 143, 265–274. <https://doi.org/10.1111/j.1744-7348.2003.tb00294.x>
- 821 Noh, H.H., Lee, J.Y., Park, H.K., Lee, J.W., Jo, S.H., Lim, J.B. et al. (2019)
822 Dissipation, persistence, and risk assessment of fluxapyroxad and penthiopyrad
823 residues in perilla leaf (*Perilla frutescens* var. *japonica* Hara). *PLoS One* 14,
824 e0212209. <https://doi.org/10.1371/journal.pone.0212209>
- 825 Omrane, S., Audéon, C., Ignace, A., Duplaix, C., Aouini, L., Kema, G., et al. (2017)
826 Plasticity of the *MFS1* Promoter Leads to Multidrug Resistance in the Wheat
827 Pathogen *Zymoseptoria tritici*. *mSphere* 2.
828 <https://doi.org/10.1128/mSphere.00393-17>
- 829 Parker, S.R., Welham, S., Paveley, N.D., Foulkes, J. & Scott, R.K. (2004) Tolerance
830 of septoria leaf blotch in winter wheat. *Plant Pathol* 53, 1–10.
831 <https://doi.org/10.1111/j.1365-3059.2004.00951.x>
- 832 Patry-Leclaire, S., Neau, E., Pitarch, A., Walker, A.-S. & Fillinger, S. (2023) Plasticity
833 of the *MFS1* promotor is not the only driver of Multidrug resistance in
834 *Zymoseptoria tritici*. *bioRxiv* 2023.12.27.573052.
835 <https://doi.org/10.1101/2023.12.27.573052>

- 836 Paveley, N., Fraaije, B., van den Bosch, F., Kildea, S., Burnett, F., Clark, W. et al.
837 (2020) Managing resistance evolving concurrently against two modes of action.
838 In: Diesing HB; Fraaije B; Mehl A; Oerke EC; Sierotzki H; Stammler G (Eds),
839 “Modern Fungicides and Antifungal Compounds”, Vol. IX, pp.141-146. Deutsche
840 Phytomedizinische Gesellschaft, Branschweig, ISBN: 978-3-941261-16-7.
- 841 Ponomarenko, A., Goodwin, S.B. & Kema, G.H.J. (2011) Septoria tritici blotch (STB)
842 of wheat. Plant Health Instructor. <https://doi.org/10.1094/PHI-I-2011-0407-01>
- 843 Rehfus, A., Strobel, D., Bryson, R. & Stammler, G. (2018) Mutations in *sdh* genes in
844 field isolates of *Zymoseptoria tritici* and impact on the sensitivity to various
845 succinate dehydrogenase inhibitors. Plant Pathol 67, 175–180.
846 <https://doi.org/10.1111/ppa.12715>
- 847 Shaw, M.W. (1990) Effects of temperature, leaf wetness and cultivar on the latent
848 period of *Mycosphaerella graminicola* on winter wheat. Plant Pathol 39, 255–
849 268. <https://doi.org/10.1111/j.1365-3059.1990.tb02501.x>
- 850 Shaw, M.W. (1989) A model of the evolution of polygenically controlled fungicide
851 resistance. Plant Pathol 38, 44–55. [https://doi.org/10.1111/j.1365-](https://doi.org/10.1111/j.1365-3059.1989.tb01426.x)
852 [3059.1989.tb01426.x](https://doi.org/10.1111/j.1365-3059.1989.tb01426.x)
- 853 Shaw, M.W. & Royle, D.J. (1989). Estimation and validation of a function describing
854 the rate at which *Mycosphaerella graminicola* causes yield loss in winter wheat.
855 Annals of Applied Biology 115, 425–442. [https://doi.org/10.1111/j.1744-](https://doi.org/10.1111/j.1744-7348.1989.tb06562.x)
856 [7348.1989.tb06562.x](https://doi.org/10.1111/j.1744-7348.1989.tb06562.x)
- 857 Suffert, F., Sache, I. & Lannou, C. (2013). Assessment of quantitative traits of
858 aggressiveness in *Mycosphaerella graminicola* on adult wheat plants. Plant
859 Pathol 62, 1330–1341. <https://doi.org/10.1111/ppa.12050>

- 860 Suffert, F., Sache, I. & Lannou, C. (2011). Early stages of septoria tritici blotch
861 epidemics of winter wheat: build-up, overseasoning, and release of primary
862 inoculum. Plant Pathol 60, 166–177. [https://doi.org/10.1111/j.1365-](https://doi.org/10.1111/j.1365-3059.2010.02369.x)
863 [3059.2010.02369.x](https://doi.org/10.1111/j.1365-3059.2010.02369.x)
- 864 Taylor, N.P. & Cunliffe, N.J. (2023a). Modelling quantitative fungicide resistance and
865 breakdown of resistant cultivars: Designing integrated disease management
866 strategies for Septoria of winter wheat. PLoS Comput Biol 19, e1010969.
867 <https://doi.org/10.1371/journal.pcbi.1010969>
- 868 Taylor, N.P. & Cunliffe, N.J. (2023b). Coupling machine learning and
869 epidemiological modelling to characterise optimal fungicide doses when
870 fungicide resistance is partial or quantitative. J R Soc Interface 20.
871 <https://doi.org/10.1098/rsif.2022.0685>
- 872 te Beest, D.E., Shaw, M.W., Pietravalle, S. & van den Bosch, F. (2009) A predictive
873 model for early-warning of Septoria leaf blotch on winter wheat. Eur J Plant
874 Pathol 124, 413–425. <https://doi.org/10.1007/s10658-009-9428-0>
- 875 The MathWorks Inc. (2022) MATLAB version: 9.13.0 (R2022b), Natick,
876 Massachusetts: The MathWorks Inc. <https://www.mathworks.com>
- 877 Torriani, S.F., Brunner, P.C., McDonald, B.A. & Sierotzki, H. (2009) QoI resistance
878 emerged independently at least 4 times in European populations of
879 *Mycosphaerella graminicola*. Pest Manag Sci 65, 155–162.
880 <https://doi.org/10.1002/ps.1662>
- 881 van den Berg, F., Paveley, N.D. & van den Bosch, F. (2016) Dose and number of
882 applications that maximize fungicide effective life exemplified by *Zymoseptoria*
883 *tritici* on wheat – a model analysis. Plant Pathol 65, 1380–1389.
884 <https://doi.org/10.1111/ppa.12558>

- 885 van den Berg, F., van den Bosch, F. & Paveley, N.D. (2013). Optimal Fungicide
886 Application Timings for Disease Control Are Also an Effective Anti-Resistance
887 Strategy: A Case Study for *Zymoseptoria tritici* (*Mycosphaerella graminicola*) on
888 Wheat. *Phytopathology* 103, 1209–1219. [https://doi.org/10.1094/PHYTO-03-13-](https://doi.org/10.1094/PHYTO-03-13-0061-R)
889 0061-R
- 890 van den Bosch, F., Oliver, R., van den Berg, F. & Paveley, N. (2014a) Governing
891 Principles Can Guide Fungicide-Resistance Management Tactics. *Annu Rev*
892 *Phytopathol* 52, 175–195. [https://doi.org/10.1146/annurev-phyto-102313-](https://doi.org/10.1146/annurev-phyto-102313-050158)
893 050158
- 894 van den Bosch, F., Paveley, N., van den Berg, F., Hobbelen, P. & Oliver, R. (2014b)
895 Mixtures as a Fungicide Resistance Management Tactic. *Phytopathology* 104,
896 1264–1273. <https://doi.org/10.1094/PHYTO-04-14-0121-RVW>
- 897 Weir, A.H., Bragg, P.L., Porter, J.R. & Rayner, J.H. (1984) A winter wheat crop
898 simulation model without water or nutrient limitations. *J Agric Sci* 102, 371–382.
899 <https://doi.org/10.1017/S0021859600042702>
- 900 Wollenberg, R.D., Donau, S.S., Taft, M.H., Balázs, Z., Giese, S., Thiel, C. et al.
901 (2020) Undeclared—Changing the phenamacril scaffold is not enough to beat
902 resistant *Fusarium*. *PLoS One* 15, e0235568.
903 <https://doi.org/10.1371/journal.pone.0235568>
- 904 Young, C., Boor, T., Corkley, I., Fraaije, B., Clark, B., Havis, N. et al. (2021)
905 Managing resistance evolving concurrently against two or more modes of action
906 to extend the effective life of fungicides. AHDB Project Report No. 637, pp 1–91.
907 Available at:
908 <https://projectblue.blob.core.windows.net/media/Default/Research%20Papers/C>

ereals%20and%20Oilseed/2021/PR637%20final%20project%20report.pdf

[Accessed: 1 August 2024].

Zadoks, J.C., Chang, T.T. & Konzak, C.F. (1974) A decimal code for the growth stages of cereals. *Weed Res* 14, 415–421. <https://doi.org/10.1111/j.1365-3180.1974.tb01084.x>

8. Supporting Information

A.1 Further details on model parameterisation

A.2 Further details on model results

9. Figure legends

FIGURE 1: Effect of asymptote shift, ζ_q , and curvature shift, ζ_k , on the dose response to fungicide dose, $D(t)$. (a) and (b) show the fractional reduction, $f(t)$, of pathogen life cycle parameters for different levels of asymptote shift and curvature shift respectively. (c) and (d) show $f_\sigma(t) - f_\rho(t)$, the resulting difference in $f(t)$ of the sensitive strain compared to a resistant strain with an asymptote shift or a curvature shift respectively. Dose response shown for a fungicide with $q_\sigma = 0.75, k_\sigma = 10$. Solid black line: dose response of sensitive strain. Dashed orange line: $\zeta_q = 50\%$. Dotted purple line: $\zeta_q = 90\%$. Solid orange line: $\zeta_k = 50\%$. Dashed purple line: $\zeta_k = 90\%$.

FIGURE 2: Effect of decay rate ν on the simulated fungicide dose, $D(t)$, and fractional reduction, $f(t)$, over time following single (solid black line) and split dose (blue dashed line) applications of a fungicide with $q = 0.75, k = 10$. (a), (b) and (c) show $D(t)$ for $\nu = 0.016 \text{ t}^{-1}$, $\nu = 0.008 \text{ t}^{-1}$ and $\nu = 0.004 \text{ t}^{-1}$ respectively, corresponding to foliar half-lives of 3, 6 and 12 days respectively. (d), (e) and (f) show $f(t)$ for $\nu = 0.016 \text{ t}^{-1}$, $\nu = 0.008 \text{ t}^{-1}$ and $\nu = 0.004 \text{ t}^{-1}$ respectively.

FIGURE 3: Model simulation of the growth, senescence and infection by *Z. tritici* of the upper wheat canopy. (a) Model simulation of healthy LAI in the absence of disease (solid line) and observed green leaf area index (GLAI) measurements used for parameterisation of wheat canopy (points) ($n=76$, from 6 sites from Dataset 1). The simulated timings of growth stages 32, 37, 39, 61 and 75 are indicated (blue arrows). (b) Model simulation of healthy (not latently infected) LAI in the presence of *Z. tritici*, latently infected LAI and infectious LAI for an average untreated epidemic of STB in the UK.

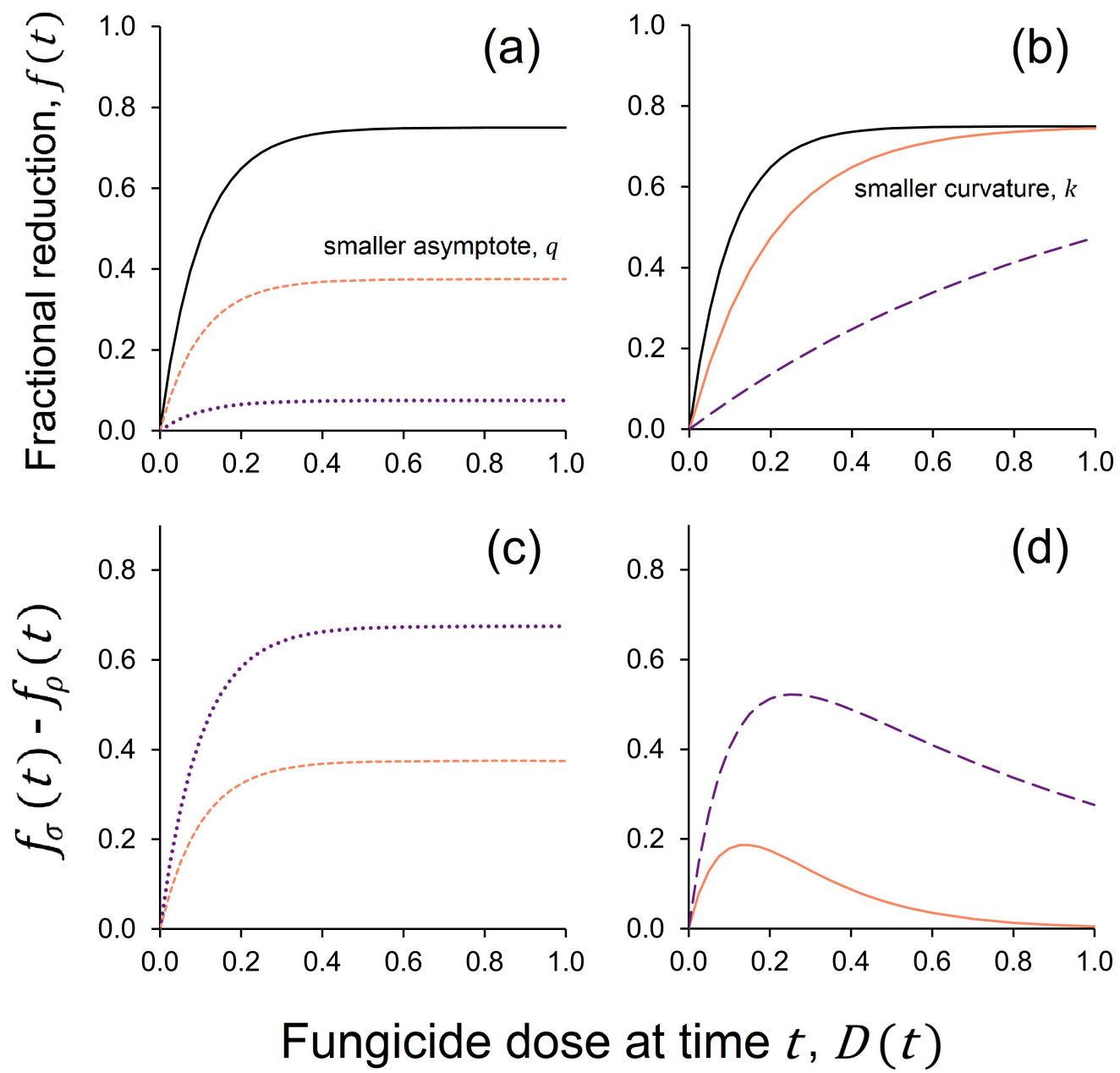
FIGURE 4: Effect of fungicide properties and resistance type on magnitude of selection for a resistant strain. Variation in selection coefficient, s with (a) asymptote parameter, q_σ ; (b) curvature parameter, k_σ ; (c) decay rate, ν ; (d) asymptote shift, ζ_q ; and (e) curvature shift, ζ_k . Only one parameter varied at a time: $\nu = 0.008$ for (a), (b), (d) and (e); $q_\sigma = 0.75$ for (b)–(e); $k_\sigma = 10$ for (a) and (c)–(e); $\zeta_q = 100\%$ for (a)–(c) and 0% for (e); $\zeta_k = 0\%$ for (a)–(d). s measures the magnitude of selection for a resistant strain.

FIGURE 5: Negligible effect of asymptote parameter, q_σ , and asymptote shift, ζ_q on η , the percentage change in selection due to dose splitting. Variation in η with (a) q_σ and (b) ζ_q for $k_\sigma = 1, 2, 5$ and 10 . (c) Variation in η with q_σ for decay rates $\nu = 0.004 \text{ t}^{-1}$, 0.008 t^{-1} and 0.016 t^{-1} . η is measured as the percentage change in selection as a result of splitting a total fungicide dose D_{Total} over two applications of $0.5D_{Max}$ at GS32 and GS39.

FIGURE 6: Percentage change in selection, η , as a result of dose splitting for a range of parameter values: curvature parameter, k_σ , decay rate, ν , and levels of sensitivity shift, ζ_q and ζ_k . Dose splitting simulated as two applications of $0.5D_{Max}$ at GS32 and

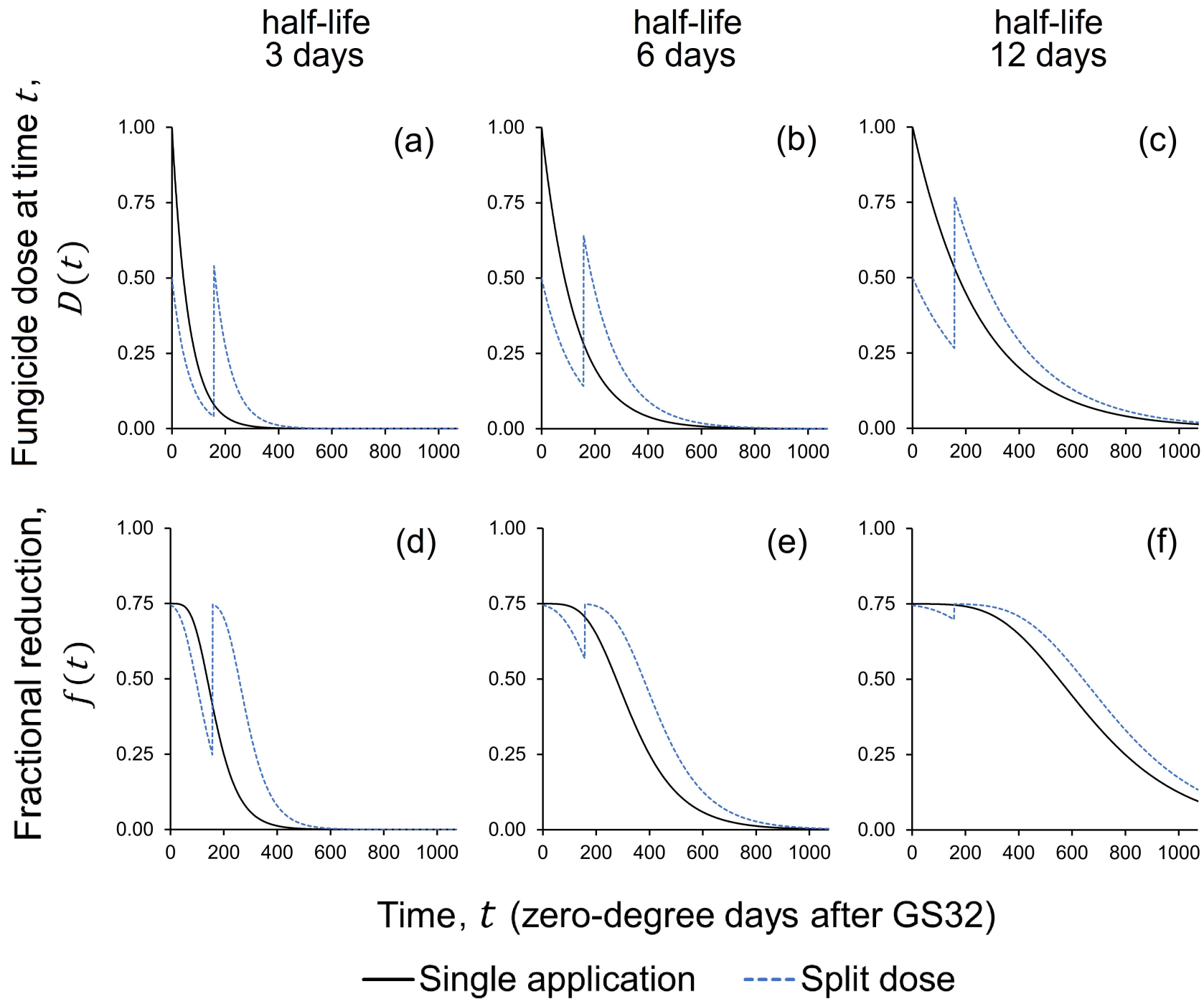
956 GS39, compared to a single application of D_{Max} at GS32. (a), (b) and (c) show the
957 effect of k_σ on η for a resistant strain with an asymptote shift, ζ_q , for fungicide decay
958 rates $\nu = 0.01605 \text{ t}^{-1}$, $\nu = 0.008 \text{ t}^{-1}$, and $\nu = 0.004 \text{ t}^{-1}$ respectively, corresponding
959 to foliar half-lives of 3, 6 and 12 days respectively. (d), (e) and (f) show the effect of k_σ
960 on η for a resistant strain with a curvature shift, ζ_k , for fungicide decay rates $\nu = 0.016$
961 t^{-1} , $\nu = 0.008 \text{ t}^{-1}$, and $\nu = 0.004 \text{ t}^{-1}$ respectively. Results shown for asymptote
962 parameter $q_\sigma = 0.5$; the effect of q_σ on η is very small (see Figure 5).

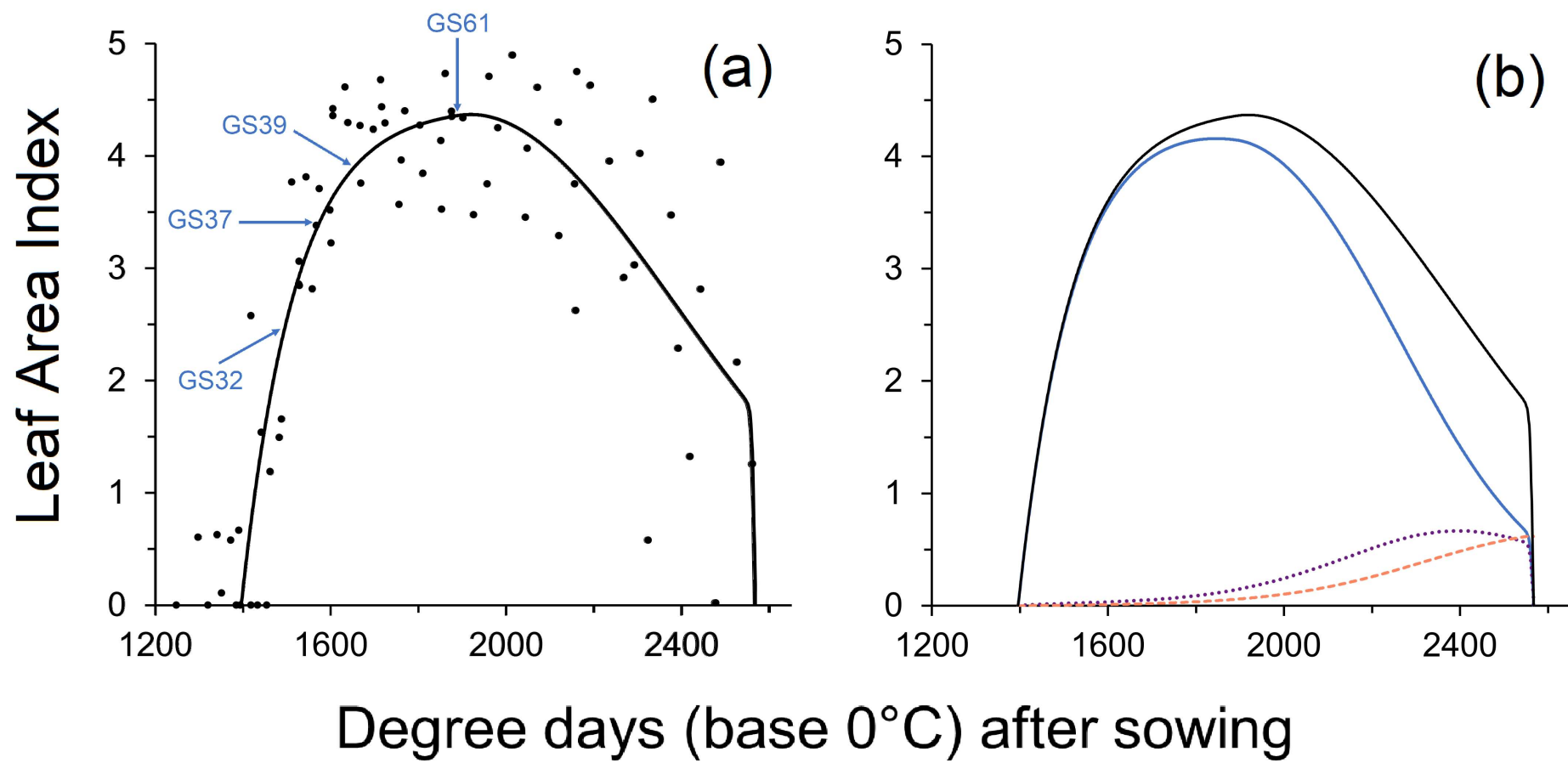
For Peer Review



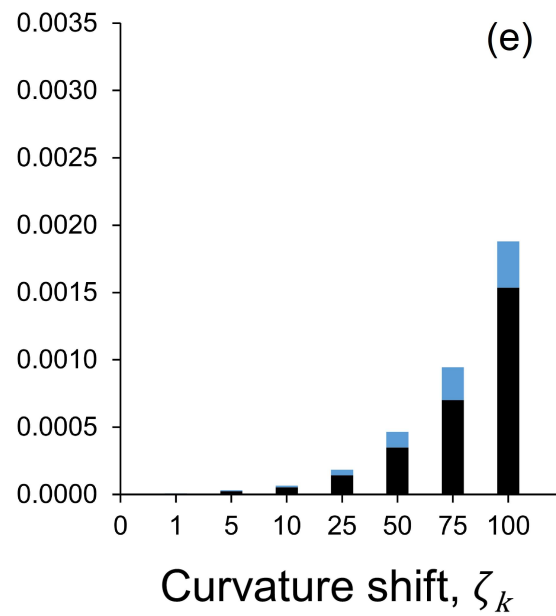
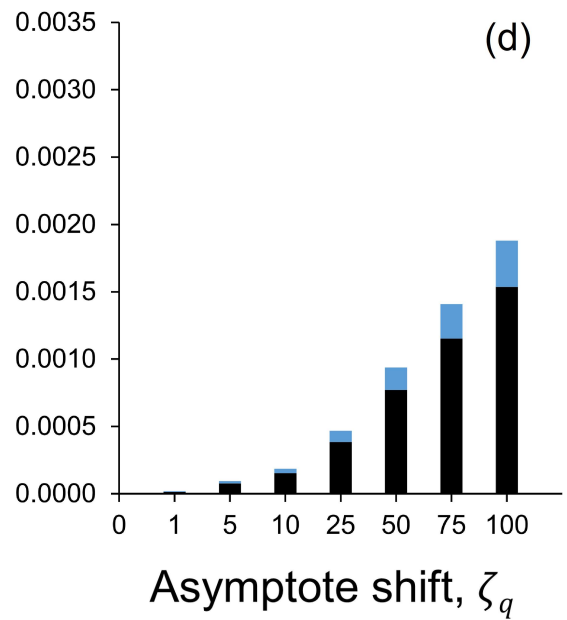
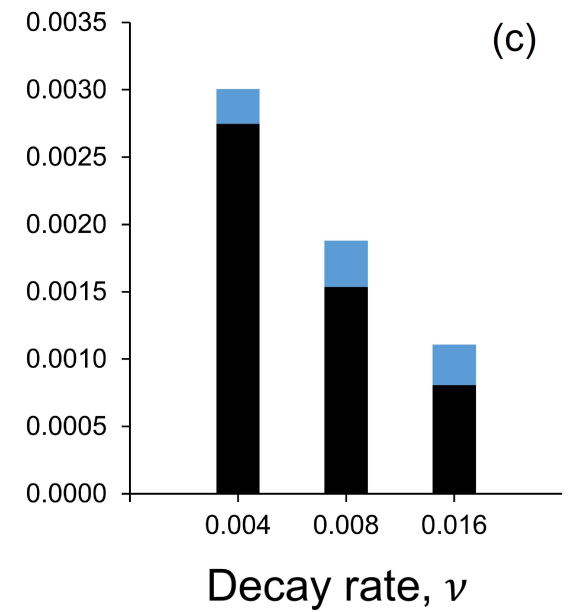
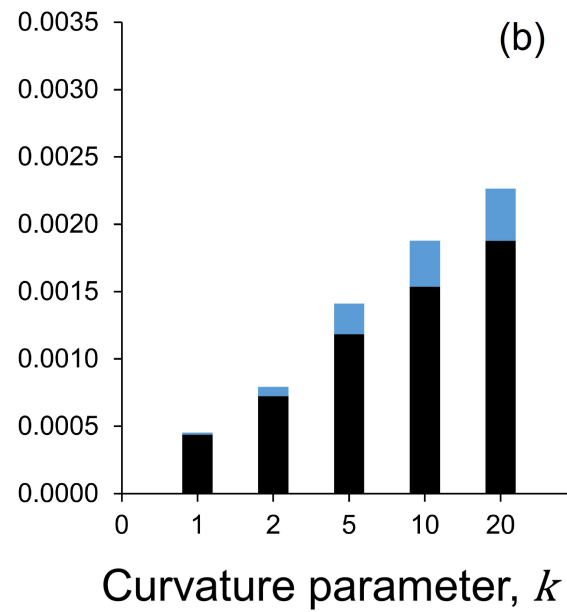
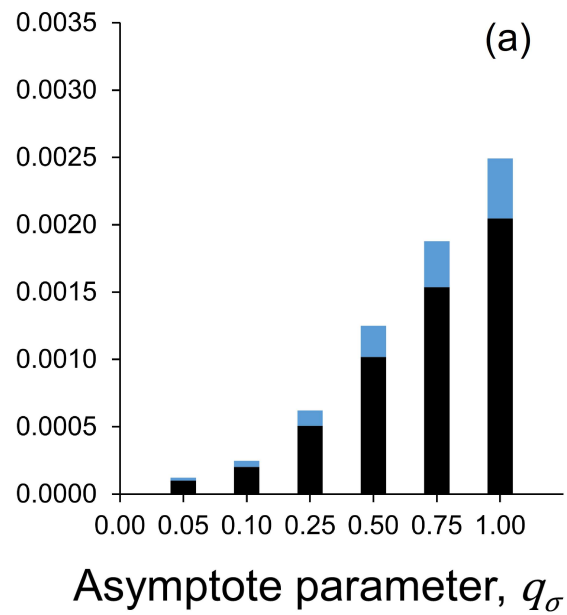
— Sensitive strain

--- 50% asymptote shift	— 50% curvature shift
... 90% asymptote shift	- - 90% curvature shift

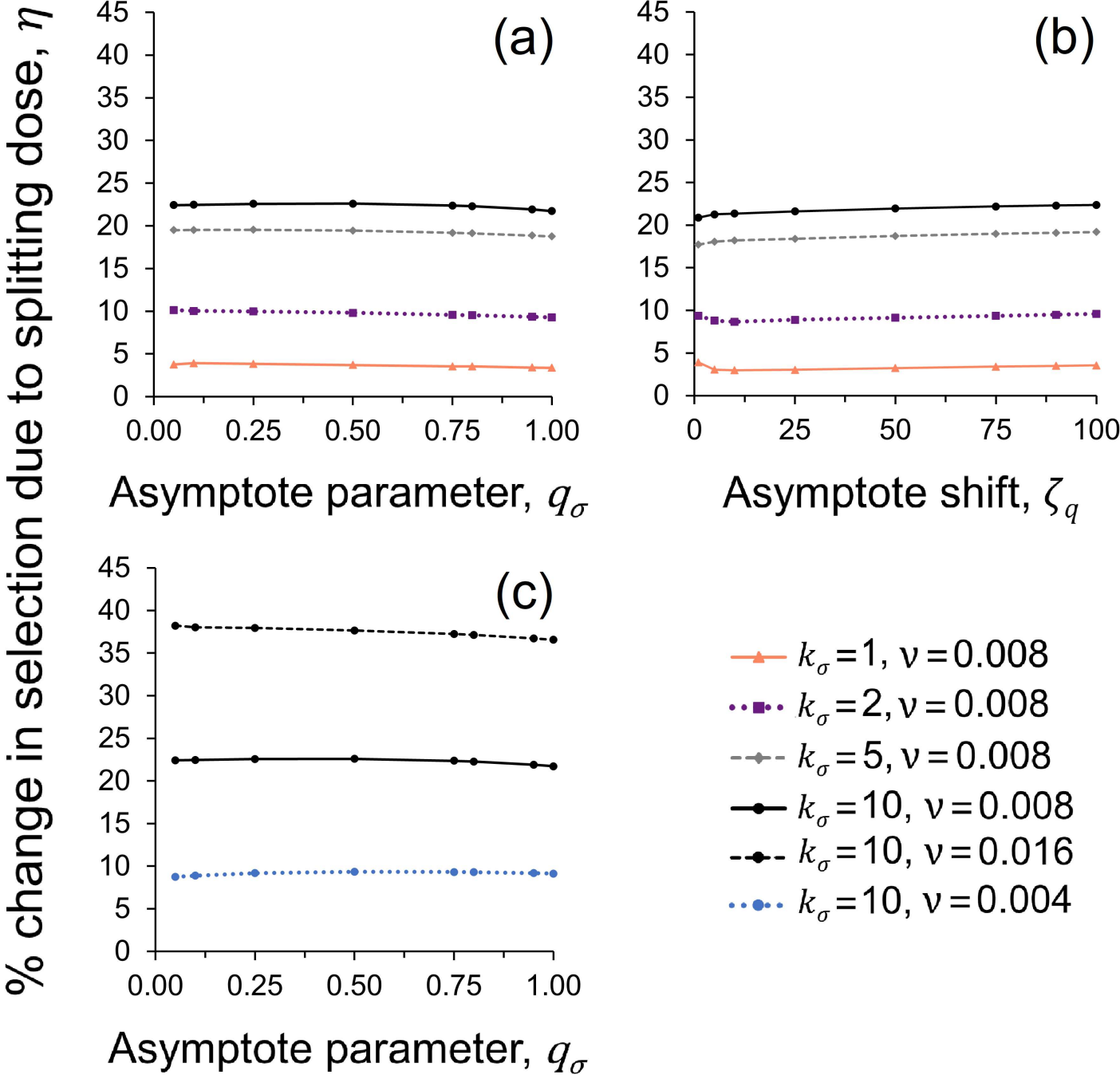




- Observed GLAI — Healthy LAI in absence of disease
- Healthy LAI in presence of *Z. tritici* Latent - - - Infectious

Selection coefficient, s 

■ Single application
■ Split dose



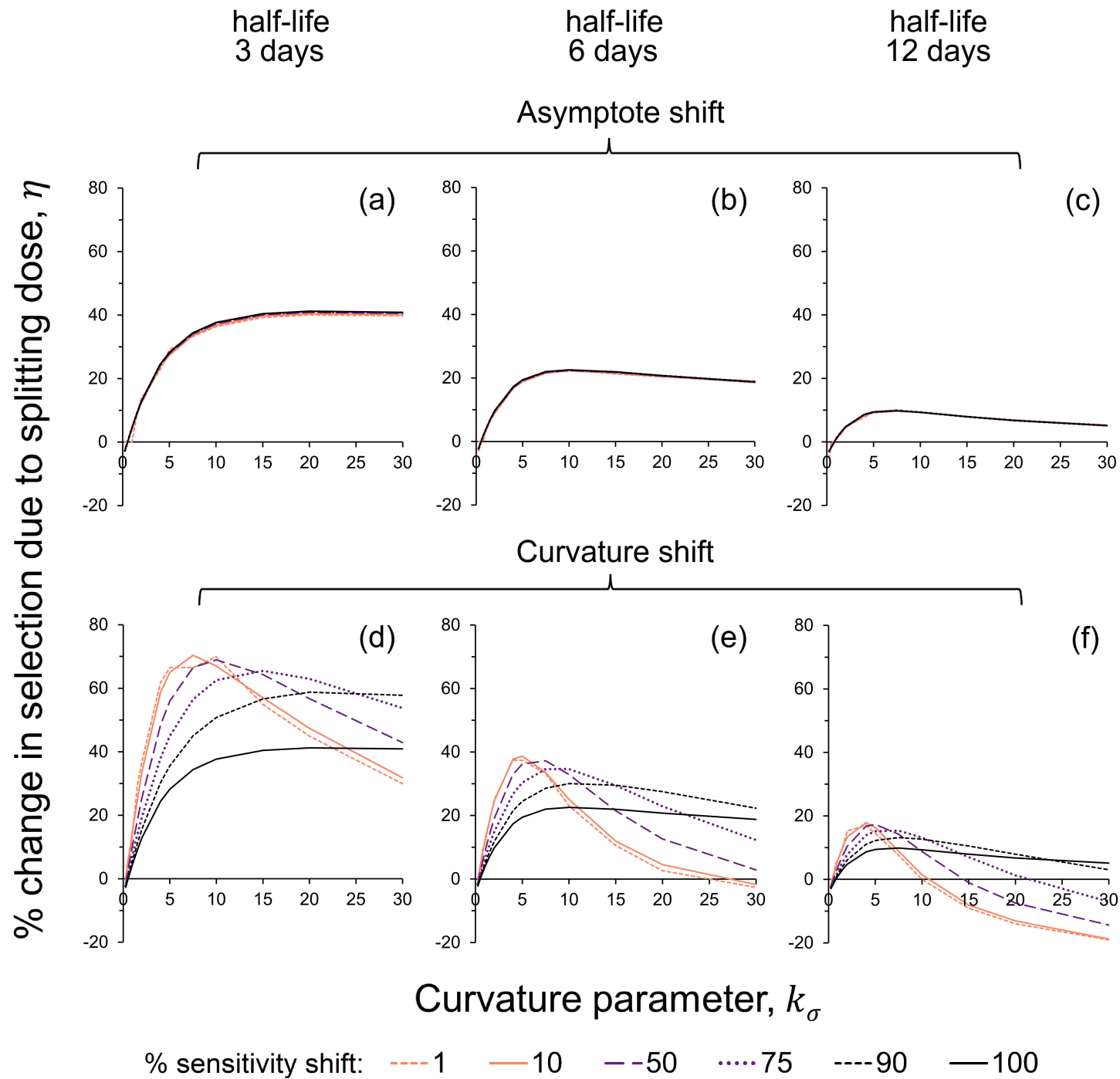


TABLE 1: List of parameter values simulated. All combinations of q_σ , k_σ and ν values simulated for each value of ζ_q and ζ_k listed.

Parameter	Description	Values simulated
D_{Total}	Total fungicide dose applied to upper leaf canopy	1, i.e. D_{Max}
$\theta_{\rho\text{Start}}$	Initial fraction of inoculum C that is resistant	0.01
q_σ	Asymptote of fungicide dose response (sensitive strain)	0.05, 0.1, 0.25, 0.5, 0.75, 0.8, 0.95, 1
k_σ	Curvature of fungicide dose response (sensitive strain)	0, 0.25, 0.5, 0.75, 1, 1.5, 2, 4, 5, 7.5, 10, 15, 20, 30
ν	Decay rate (t^{-1})	0.01605, 0.00802, 0.00401
ζ_q	Asymptote shift of resistant strain	0, 1, 5, 10, 25, 50, 75, 90, 100
ζ_k	Curvature shift of resistant strain	0, 1, 5, 10, 25, 50, 75, 90, 100
GS32	Timing of GS32 application (zero-degree days)	1495
GS39	Timing of GS39 applications (zero-degree days)	1653

TABLE 2: Fitted parameter values. Time, t is measured in degree days (base 0°C) after sowing. ^aEstimate based on 'Data set 1' from Milne et al., 2003; ^bShaw, 1990; Suffert et al., 2013; ^cBoixel, 2020; Eyal, 1971; ^dHobbelen et al., 2011b; ^eEstimate based on data from AHDB Fungicide Performance field trials; ^fFantke et al., 2014; He et al., 2016; Noh et al., 2019.

Parameter	Definition	Units	Fitted value	Source
t_0 , GS31	Timing of start of growth of leaf 3	t	1396	a
GS32	Timing of GS32: leaf 3 fully emerged	t	1495	a
GS37	Timing of GS37: leaf 2 fully emerged	t	1574	a
GS39	Timing of GS39: flag leaf fully emerged	t	1653	a
t_{β_0} , GS61	Timing of anthesis & start of leaf 3 senescence	t	1891	a
t_{β_T} , GS87	Timing of end of grainfill & complete senescence of wheat canopy	t	2567	a
A_{Max}	Maximum leaf area index of top three leaves of the wheat canopy	-	4.438	a
γ	Growth rate of leaf area	t^{-1}	0.0082	a
τ	Coefficients controlling the rate of senescence over time, in relation to the length of time after the onset of senescence	t^{-1}	0.0028	a
φ		t^{-1}	0.704	a
ω		t^{-1}	0.314	a
$1/\delta$	Average latent period	t	350	b
$1/\mu$	Average infectious period	t	600	c
C_0	Initial density of infectious lesions on the lower leaves	-	0.0144	a
λ	Rate at which $C(t)$ decreases	t^{-1}	0.00897	d
ε_0	Transmission rate	-	0.0211	a
z	Number of zero-degree days per day	t	14.4	a
q_σ	Asymptote parameter for an SDHI fungicide (against sensitive strain)	-	0.569	e
k_σ	Curvature parameter for an SDHI fungicide (against sensitive strain)	-	9.9	e
ν	Decay rate for an SDHI fungicide	t^{-1}	0.00802	f